

The Effect of Functional Diversification on the Credit Risk of Financial Institutions

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Jani Nurmi
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Author	Jani Nurmi	
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Abstract**PURPOSE AND MOTIVATION:**

The purpose of this study is to examine the effect functional diversification on the credit risk of financial institutions. The focus is on whether financial institutions should be allowed to combine commercial banking and investment banking functions under one financial conglomerate. Under current regulators financial institutions are allowed to combine the functions, however regulators are considering driving regulation towards separating commercial and investment banking. This study sheds light on whether the contemplated separation of investment and commercial banking increases or decreases the credit risk of financial institutions.

DATA AND METHODS:

The data sample of his study consists of 51 financial institutions from Europe and the US, with a time span ranging from 2007 to 2014. Functional diversification is measured using income and asset based measures, collected from the annual financial statements of the financial institutions. Credit risk is measured with the financial institution level credit default swap (CDS) spreads and the CDS spread difference with banking sector CDS index spreads. The impact of functional diversification on credit risk is examined with multiple panel data regressions, where the credit risk of financial institutions is explained with the functional diversification measures and a set of control variables. Furthermore, multiple robustness checks are developed.

FINDINGS:

Based on the results, functional diversification decreases the credit risk of financial institutions at the financial institution credit risk level and compared with the average credit risk in the banking sector. The results are confirmed with multiple robustness checks. Based on the results the contemplated separation of commercial and investment banking with new regulation can have grave consequences. Separating the two functions would decrease functional diversification increasing the credit risk of financial institutions and the probability of financial institution failures.

Keywords Financial institutions, functional diversification, credit risk, credit default swaps, economies of scope, agency costs, conflicts of interest, regulation of financial institution

Tekijä Jani Nurmi

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Tiivistelmä**TUTKIMUKSEN TARKOITUS:**

Tämän tutkimuksen tarkoitus on tarkastella rahoituslaitosten funktionaalisen hajauttamisen vaikutusta laitosten luottorisktiin. Fokuksena on tutkia, pitäisikö rahoituslaitoksille antaa mahdollisuus yhdistää investointipankki- ja liikepankkifunktiot yhden rahoituslaitoksen yhteyteen. Nykyisen regulaation mukaan rahoituslaitokset saavat yhdistää funktiot. Tästä huolimatta rahoituslaitosten valvojat tällä hetkellä harkitsevat regulaation muuttamista suuntaan, jossa investointipankki- ja liikepankkifunktiot erotetaan toisistaan. Tämä tutkimus analysoi, kasvattaako vai laskeeko mahdollinen uusi regulaatio rahoituslaitosten luottoriskiä.

AINEISTO JA MENETELMÄT:

Tutkimuksessa käytetty dataotos koostuu 51 rahoituslaitoksesta Euroopasta ja Yhdysvalloista. Otoksen aikajakso alkaa vuodesta 2007 ja loppuu vuoteen 2014. Funktionaalista hajautusta mitataan käyttämällä liikevaihto- ja taseperusteisia mittareita, jotka on kerätty rahoituslaitosten vuosittaisista tilinpäätöksistä. Luottoriskiä mitataan käyttämällä rahoituslaitosten luottotappioriskin vaihtosopimusten (credit default swap, CDS) spredejä ja spredien eroa pankkisektorin CDS-indekseihin. Funktionaalisen hajautuksen vaikutusta luottorisktiin tutkitaan useilla paneelidataregressioilla, joissa rahoituslaitosten luottoriskiä selitetään funktionaalisen hajautuksen mittareilla ja monilla kontrollimuuttujilla. Lisäksi useita lisätestejä käytetään tulosten vahvistamiseksi.

TULOKSET:

Tulosten perusteella funktionaalinen hajautus vähentää rahoituslaitosten luottoriskiä rahoituslaitostasolla sekä verrattuna keskimääräiseen pankkisektorin luottoriskitasoon. Tulokset vahvistetaan useissa lisätesteissä. Tulosten mukaan suunnitteilla olevalla investointipankki- ja liikepankkifunktioiden erottavalla uudella regulaatiolla voi olla merkittävä vaikutus rahoituslaitosten luottorisktiin. Uusi regulaatio laskisi funktionaalisen hajautuksen tasoa, joka puolestaan nostaisi rahoituslaitosten luottoriskiä ja niiden konkurssin todennäköisyyttä.

Avainsanat Rahoituslaitokset, funktionaalinen hajautus, luottoriski, luottotappioriskin vaihtosopimukset, laajuuden ekonomia, agenttikustannukset, intressiristiriidat, rahoituslaitosten regulaatio

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1. Introduction & Motivation

The impact of functional diversification (i.e. combining commercial banking, investment banking and other financial services under the same financial conglomerate) on the riskiness and value of financial institutions has been subject to much debate among academics and regulators. Current literature offers conflicting results on the impact of financial institution functional diversification, some suggest higher diversification decreases the riskiness of financial institutions and some the opposite. According to the standard portfolio theory (Markowitz 1952) diversification reduces the volatility and riskiness of a portfolio of assets. According to Demsetz & Trahan (1997) and Baele, De Jonghe & Vander Vennet (2007) when applying the portfolio theory to banking and financial services, functional diversification should decrease the probability of default and riskiness of the financial institution. However, some suggest that the increasing amount of different financial services offered by the financial institution increases agency costs and conflicts of interest inside the financial institution, and thus its riskiness.

Regulation currently allows financial institutions to combine traditional commercial banking with investment banking under one financial conglomerate. However, historically commercial banking and investment banking has not been allowed to operate under the same legal entity. In the US, the Glass Stegall act was introduced in 1933 as a response to the great depression, prohibiting financial institutions to combine investments banking services with commercial banking and limited the ability of commercial banks to offer securities activities, such as securities underwriting. In 1999, the Gramm–Leach–Bliley Act was introduced, allowing financial institutions to functionally diversify and combine commercial and investment banking under Financial Holding Companies. In the European Union, the Second Banking Directive of 1989 allowed European financial institutions to functionally diversify across different financial services, allowing financial institutions to offer commercial banking, investment banking, insurance and other financial services through the same legal entity. New regulation in Europe and the US has allowed financial institutions to pursue a wide range of diversification strategies, some financial institutions have remained focused on the traditional commercial banking market or on investment banking, and others have pursued to offer a wide range of financial services becoming large financial conglomerates.

After the financial crisis, regulators and politicians in the US and EU have once again suggested to separate commercial banking from investment banking. In the EU, the Liikanen report calls for the separation between high risk trading from commercial banking services. The report does not suggest the total legal separation of high risk trading and commercial banking under one financial conglomerate. However, activities would have to be conducted under different entities owned by the conglomerate to avoid agency costs and conflicts of interest. The FED is contemplating making it so costly and difficult for functionally diversified financial institution to operate that they eventually divide themselves up, effectively re-enforcing Glass Stegall¹. Similarly, the bank of England and BaFin (the German financial institution regulatory body) hint that they are examining the possibility to introduce new regulation, separating commercial and investment banking². Furthermore, major global financial institutions are considering whether to split up commercial and investment banking, Deutsche Bank is considering spinning off its consumer banking division and analysts are pressuring JPMorgan to split its commercial and investment banking divisions¹.

Since the aim of the potential new regulation is to decrease the functional diversification of financial institutions, the motivation of this study is to find out how functional diversification impacts the credit risk (i.e. probability of default) of financial institutions. Previous studies have mainly focused on measuring the impact of functional diversification on the equity market risk of financial institutions. The equity risk measures, such as beta, market value, share price volatility, z-scores and Tobin's q have been used to determine the relationship between risk and functional diversification. However, the equity risk measures capture a large number of risk factors that impact the results. For regulatory purposes and to determine whether commercial and investment banking should be allowed to operate under one financial institution, the most appropriate risk measure is the credit risk of the financial institution. The purpose of financial institution regulation is to mostly protect depositors and the overall economy from the adverse effects of defaulting and distressed financial institutions. Financial institutions offer multiple financial services that are crucial to a well-functioning economy, acting as intermediaries between depositors and corporates in need of financing and providing payment services which are crucial in a modern economy. Defaults of financial institutions can have very adverse effect on the economy as deposits can be at risk (even though deposits are mostly guaranteed by

¹ Sources: Financial Times, "Regulators right to cut biggest banks down to size", 7.1.2015 & "Regulators test the universal banking model", 15.1.2015

² Source: The Economist, "Together, forever? The enduring marriage of investment and commercial banking", 18.8.2012

governments) and money transfers can slow down in the case of disturbance in the banking sector. Financial institution regulation is mostly in place to avoid financial institution defaults and the adverse effect resulting from the defaults, making credit risk a major point of interest from the regulative perspective³. Thus in this study, instead of measuring financial institution level risk using risk measures from the equity market, credit default swap (CDS) spreads of the financial institutions are used. CDS yields directly measure the credit risk (i.e. probability of default) of financial institutions. The results should provide insights from a regulatory perspective on whether functional diversification increases or decreases the credit risk of financial institutions and what is the impact of separating investment banking from commercial banking on the credit risk of financial institutions.

The data sample used in this study consists of 51 financial institutions from the US and Europe. The sample is Europe focused with 9 financial institutions from the US, since only 9 financial institutions from the US have liquid CDSs traded. The time span of the sample starts from 2007 and ends at the end of 2014. The financial crisis started during the time span of the sample and had a major effect on the 2008 observations. However, multiple robustness checks have been performed in order to mitigate the effect of the financial crisis.

In this study functional diversification is measured with the income and asset diversification of financial institutions. Income diversification measures the extent to which the income mix between net interest income (generated by traditional commercial banking activities) and non-interest income (mostly fee and trading related income generated from investment banking) of the financial institution is diversified. Even though non-interest income can include income from insurance and other financial services, the majority of non-interest income is investment banking related. Thus financial institutions generating more non-interest income (interest income) are referred as investment banks (commercial banks). Asset diversification refers to the extent to which the mix between the loans (commercial banking assets) and other earning assets (investment banking assets) in the balance sheet of the financial institution is diversified.

The main focus of this study is to examine how functional diversification impacts the credit risk of financial institutions. However, at first it is examined whether financial institutions

³ Information about the purpose of financial institution regulation has been obtained from Division of Supervision and Risk Management of Federal Reserve Bank of Kansas City, "Banking Regulation", 2000

that focus more on investment banking⁴ have higher credit risk than those focusing more on commercial banking. The impact of functional focus needs to be determined first since diversification benefits could arise from less exposure to a riskier function and not functional diversification itself. When actually examining the impact of functional diversification the functional focus is controlled for. Based on the results, the functional focus of financial institutions does not seem to impact their credit risk. Financial institutions focusing more on investment banking do not differ in terms of credit risk from financial institutions focusing more on commercial banking.

When examining functional diversification, it is found that financial institutions with higher income diversification have lower CDS spreads at the financial institution level and compared with the banking sector CDS index. The results suggest that benefits from functional diversification decrease the credit risk of financial institutions and should make them less probable to default. I find no evidence that asset based diversification impacts the credit risk of financial institutions. However, income based measures are more appropriate, they better describe the functional focus of the financial institution since asset diversification measures do not fully capture the extent of investment banking services. Based on the results, the current regulatory agenda to decrease the level of functional diversification of financial institutions, seems counterintuitive. If new regulation is introduced, separating commercial and investment banking would increase credit risk, making financial institutions more likely to default. The motivation for the new regulation seems to stem from the desire to make monitoring of large and complex functionally diversified financial institutions easier. However, the benefits from easier regulation and monitoring should not outweigh the grave consequences of increased probability for financial institution failures.

The rest of this study is constructed as follows. Chapter 2 describes previous literature focusing on the functional diversification of financial institutions. Chapter 3 describes the data and methods used in this study. Chapter 4 outlines the hypothesis. Chapter 5 describes the results in detail. In Chapter 6 the robustness of the results is tested, chapter 7 discusses the limitations of the study and chapter 8 concludes.

⁴ For the entirety of this study financial institution focus areas are divided into commercial banking and investment banking. Even though when measuring the focus non-interest income can contain income from non-investment banking functions (e.g. insurance or other financial services). However, since investment banking dominates the non-interest income, only investment banking/commercial banking division is used.

2. Literary Review

More recent literature examining the functional diversification of financial institutions does not focus on studying the economies of scope or conflicts of interest related to functional diversification directly, but the overall impact of functional diversification on the riskiness and value of financial institutions. A similar approach is used in this study and the aim of this study is to contribute to the literature examining the overall impact of functional diversification. The academic results vary between different academics, some papers find that functional diversification increases (decreases) risk (value) and others the opposite. Broadly research on European financial institutions differs to the results found in the US and other developed markets. Older research has tried to directly examine the impact of conflicts of interest or economies of scope resulting from functional diversification, without achieving consensus. The newer studies focusing on measuring the overall impact of functional diversification, directly contribute to the older research. If financial institutions with high diversification are found to have higher (lower) risk (value), the existence of conflicts of interest and agency costs could explain the discount. Vice versa lower (higher) risk (value) could suggest the existence of diversification benefits or economies of scope.

At first in the literature review, the more recent research focusing on the overall impact of functional diversification is discussed. Secondly older research on conflicts of interest and economies of scope is discussed. Finally research on why CDSs should be used to measure credit risk is discussed.

2.1. The Overall Impact of Functional Diversification on the Riskiness and Value of Financial Institutions

This section is divided into two sub-sections, first one describing studies that find functional diversification increasing the riskiness of financial institutions and the second detailing papers discovering the opposite.

2.1.1. Research Suggesting Functional Diversification Increases the Riskiness of Financial Institutions

Based on US financial institutions Stiroh (2006) suggests that increased share of non-interest income, income mostly from investment banking, raises the volatility of profits without raising the average profit of the financial institutions. Moreover, Stiroh (2006) finds a positive relationship between non-interest income and the volatility of total and idiosyncratic risk components. He also finds a positive relationship between the market betas and a higher share of non-interest income of financial institutions. Finally Stiroh (2006) suggests, that some US financial institutions with high non-interest income shares have over diversified their activities, since their riskiness have increased without increase in average profitability.

Using cross-sectional analysis Stiroh and Rumble (2006) find that more functionally diversified financial institutions in the US have higher risk-adjusted profits, however the effect of higher profits is offset by costs associated to higher exposure to more volatile non-interest activities. They conclude that diversification benefits can be gained, but high volatility activities like trading, diminish the diversification benefits on the risk adjusted basis. Furthermore, using a panel data sample they confirm that higher share of non-interest activities is negatively associated with risk-adjusted profits and find no evidence that diversification affects the performance or profitability of US financial holding companies. They conclude that financial holding companies have overestimated the benefits of functional diversification and argue that even though functional diversification may generate new income opportunities for financial institutions, the different income stream are subject to the same industry shocks and risks reducing diversification benefits.

Using a sample of financial institutions from 43 developed countries Laeven and Levine (2007) examine whether the Tobin's q of financial institutions is higher than if the financial institution were to be separated to single corporations, where each entity would specialize in offering a specific financial service. According to them, the existence economies of scope would result in higher valuations of diversified financial institutions compared with if the financial institutions were broken into separated specialized entities. Vice versa the existence of agency costs would result in a valuation discount on the diversified financial institutions. They find that market values of diversified financial institutions are significantly lower than if the financial institutions were to be broken into specialized independent entities. They argue

that diversification of financial institutions increases agency costs and has a higher impact than any potential economies of scope. Thus, if financial conglomerates have lower valuations than if the same conglomerate was operating as specialized individual entities, functional diversification should increase the riskiness of financial institutions. Using a sample of US financial institutions Schmid and Walter (2009) find similar results, diversified financial conglomerates have lower market valuations. The lower market value is evident when financial institutions diversify with commercial banking, insurance and securities services, no impact in value is seen in financial institutions operating mainly in investment banking activities (Schmid, Walter 2009). They speculate that the reduction in value of diversified financial institutions is caused by the overestimations of the potential economies of scope related to diversification.

2.1.2. Research Suggesting Functional Diversification Decreases the Riskiness of Financial Institutions

Examining listed financial institutions from 17 European countries with a panel data sample covering 1989-2004, Baele, De Jonghe and Vander Venet (2007) find a positive relationship with functional diversification and the franchise value of financial institutions, suggesting that the equity markets value functionally diversified financial institutions higher than less diversified counterparts. They find that the more functionally diversified financial institutions have higher market risk and lower idiosyncratic risk. Thus more functionally diversified financial institutions have higher market betas and higher systematic risk and lower financial institution specific risk. Financial institutions that diversify become more similar to the average financial institution in the market, explaining the higher market betas. Lower idiosyncratic risk suggests that functional diversification decreases the financial institution specific risk, reducing their probability of default and thus credit risk.

Using a sample of financial institutions from 9 developed countries, including the US, Elsas, Hackethal and Holzhäuser (2010) find that financial institutions diversification increases profitability through higher margins from non-interest services and lower cost to income ratios. According to them, higher profitability increases the market valuations of diversified financial institutions. When, they control for profitability they find no impact of diversification on valuations, implying limited evidence for the existence of conflicts of interest and that economies of scope are evident in diversification. The authors suggest that previous results finding a negative diversification impact on valuations are driven by the use of insufficient

measures for diversification and market valuations, and not using control variables for profitability.

Demirgüç-Kunt and Huizinga (2010) examine the effect of diversification and short-term funding strategies of financial institutions from 101 countries on their risk and return, mainly the rate of return on assets and distance to default. They discover that expansion into non-interest income generating services can offer some diversification benefits at low levels of non-interest income. They conclude that financial institutions relying mostly on non-interest income and on non-deposit funding (mostly specialized investment banks) are very risky. In their opinion the downfall of the US investment banking sector during the financial crisis can be explained on their over-reliance on risky non-interest income and non-deposit short term funding. They conclude that evidence on diversification benefits is weak, there can be some benefits on combining risky investment banking with traditional commercial banking. However, financial institutions should not overly depend on non-interest income generating activities.

De Jonghe (2010) examines the relationship between financial institution diversification and the ability of the financial institution to survive a banking sector crash. He finds that non-interest income focus increases the tail-beta of financial institutions, reducing banking system stability. He finds that smaller better capitalized banks focusing mainly on commercial banking, are more likely to sustain difficult market conditions and improving the overall banking sector stability. Wagner (2010) shows, using models, that diversification reduces the probability of default of individual financial institutions. However, he finds that diversification reduces the stability of the overall banking sector increasing the likelihood of systemic crises, since diversification makes financial institutions more similar and more correlated.

The literature on the impact of functional diversification on the riskiness and value of financial institutions is mixed. Broadly, European evidence suggests that diversification decreases risk and evidence from US the opposite. However, even geographical consensus cannot be established. This study aims to build on previous research, by introducing an alternative perspective using credit risk as the explanatory variable. Credit risk is better suited to determine whether functional diversification should be allowed from the regulatory point of view. Previous studies have mainly examined the impact of functional diversification from the shareholder perspective, using equity based measures for risk and value. Based on the perspective different conclusions can be drawn. It can be possible to transfer value and risk between the share- and debtholders. According to Green and Talmor (1986) and Jensen and

Meckling (1976) the asset substitution problem can cause value transfers between the share- and debtholders of a corporation. According to them if a company is financed with high amounts of debt, shareholders might increase the riskiness of the company's projects in the expense of the debtholders, transferring risk to and value from the debtholders. Thus results on the impact of functional diversification on risk, can be different in the debt and equity markets. Making regulatory decisions based on studies using the shareholder perspective can lead to wrong conclusions, as the credit risk perspective is better suited for regulatory conclusions. Since mixed results has been obtained using equity data, the credit risk perspective should shed more light on whether functional diversification should be allowed and commercial banking combined with investment banking. However, it can't be said that result obtained from the CDS market can be generalized to the equity market. Potential risk transfer in financial institution diversification is not the focus of this study but should provide an interesting topic for further research, as the results from the equity market seem inconclusive.

2.2. Economies of Scope and Conflicts of Interest

The older literature studying the impact of functional diversification aims to directly identify and discover the potential economies of scope and conflicts of interest associated with functional diversification. Unlike the more recent literature, the older studies do not directly measure how functional diversification impacts the riskiness of the whole financial institution. The older literature does not directly compare with this study, but provides the theoretical and empirical backbone, on how functional diversification can create the economies of scope and conflicts of interest and through them impact the riskiness of the financial institution.

According to Baele, De Jonghe and Vander Vennet (2007) the potential advantages of functionally diversified financial institutions result from the potential enhanced revenue generation ability of the financial conglomerate or potential synergies. They argue that the potential synergies might result from the increased economies of scope through the sharing of employees, information and technology between different services offered by the financial conglomerate.

The potential economies of scope are argued to result from the sharing of information within the financial conglomerate. Financial institutions benefit from the information acquired from the customers they lend to in traditional commercial banking, they receive private information from the lending relationship unavailable to the public. The information can be

utilized when offering other financial services, such as investment banking or insurance services to the same customers, improving the provision of financial services (Diamond 1991, Rajan 1992, Stein 2002). Vice versa, other financial services such as securities underwriting, insurance, mutual fund and brokerage services generate information that can be used in the loan origination in commercial banking and improve the way financial institutions originate loans (Diamond 1991, Rajan 1992, Stein 2002). Since financial conglomerates can utilize the information acquired from a single financial service in other services, financial conglomerates should benefit from the economies of scope which should improve performance and reduce credit risk.

According to Vander Venet (2002) functionally diversified financial institutions are not more cost efficient, compared with specialized financial institutions, in producing traditional commercial banking services. However, Vander Venet (2002) finds that diversified financial institutions are more cost efficient when other investment banking services are taken into account. They also discover that diversified financial institutions are more efficient in terms of operational efficiency and profit efficiency.

Using a framework which takes into account the potential revenue and cost synergies, DeYoung and Roland (2001) find when financial institutions move their service mix from traditional commercial banking towards fee based investment banking, the revenue volatility and leverage of the financial institution increases, suggesting an increase in overall earnings volatility. Furthermore, they find that the overall earnings level also increases when the service mix is moved towards investment banking, suggesting a risk premium on investment banking. Stiroh (2004) finds similar evidence suggesting small or no diversification benefits from shifting the service mix towards fee income generating services. Stiroh (2004) finds a negative relationship between the non-interest income share and profit per unit of risk, especially trading activities appear to decrease profit per unit of risk. Stiroh and Rumble (2006) conclude that some diversification benefits exist for US financial holding companies, however these benefits are offset by increases in more volatile investment banking which is not more profitable than traditional interest generating commercial banking services.

Mercieca, Schaeck and Wolfe (2007) find that small European financial institutions do not benefit from diversification. They find non-interest income services negatively related to profitability and risk-adjusted performance. They conclude that small European financial institutions do not show the existence of economies of scope through diversification.

Measuring the economies of scope related to the functional diversification of financial institutions has proven difficult. According to Berger and Humphrey (1997) the identification of cost functions in diversified financial conglomerates offering a wide range of financial services, suffer from econometric challenges. Berger and Humphrey (1997) suggest that the econometric difficulties explain to some extent why the existence of economies of scope has not been confirmed or rejected in offering diversified financial services.

Overall, according to previous research it seems that offering investment banking together with commercial banking does not seem to increase the profitability of financial institutions on risk adjusted basis. Offering limited evidence on the existence of economies of scope in financial services.

Functional diversification may create or intensify agency costs and conflicts of interest inside the financial institution, between the insiders and outsiders and between customers and the financial institution. The existence of conflicts of interest when a financial institution originates loans in commercial banking and offers investment banking services, mostly securities underwriting and distributing, to the same corporate customer has been extensively examined. The literature examines whether financial institutions, which lend to a corporate client and underwrite its securities, try to generate private benefits by selling the securities of the client corporation to the public as higher quality assets than they actually are.

According to Kroszner and Rajan (1994) financial institutions lending to a corporate client receive information about the client before the public, this may incline the financial institution to underwrite the securities of the client to the unaware public at inflated prices. The purpose is to subsidize the lending through underwriting inflated securities. They argue that since financial institutions have access to a large number of unsophisticated depositors, the financial institutions have the opportunity to distribute the inflated lower quality securities without the public's full understanding of the true quality of the securities. However, by examining US financial institutions before the introduction of the commercial and investment banking separating Glass-Steagall Act of 1933, Kroszner and Rajan (1994) find that investors and rating agencies took into account the possibility of conflicts of interest when the same financial institution both underwrote the securities and had a lending relationship with company in question. They found that financial institutions with underwriting and lending relationships were forced to mainly underwrite better quality securities, since the public took into account the possible conflicts of interest in lower quality and more information sensitive securities.

Moreover, Puri (1996) finds that investors are willing to pay a higher price for securities underwritten by financial institutions with lending relationships, since the financial institution possesses private information. Gande et al. (1997) finds similar evidence; investors pay higher prices (lower yields), for low grade debt securities underwritten by financial conglomerates that have a lending relationship with the company in question. Gande et al. (1997) find no evidence of conflicts of interest even when a debt issue (underwritten by a financial institution with debt to the issuing company) is used to refinance the debt owed to the underwriting financial institution, a situation where the potential for conflicts of interest is high.

More recent research find no evidence supporting that debt underwritten by a financial institution with a lending relationship with the debt issuing corporation has lower yields (Sufi 2004), suggesting that financial institutions with lending relationships do not have an information advantage over non lending relationship financial institutions. Furthermore, Sufi (2004) finds that debt underwritten by lending relationship financial institutions has lower underwriting fees and confirms that the lower fees are not offered to capture business in the future. The lower underwriting fees are confirmed by Drucker and Puri (2005) and they find lower yield spreads for corporations whose debt is underwritten by a financial institution with a lending relationship to the same corporation, suggesting the existence of economies of scope and that the customers of the financial institution also benefit from them.

Schenone (2004) examines whether a lending relationship with a financial institution affects the underpricing of corporate IPO's, when the same financial institution underwrites the corporations securities. She finds that a pre-IPO lending relationship reduces underpricing. When underwriting securities financial institutions can reduce the asymmetric information gap between the public and the issuer, by utilizing the private information gained by the pre-IPO lending relationship. However, Ber, Yafeh and Yosha (2001) find conflicting evidence from the Israeli IPO market, they find that companies with IPO's underwritten by a financial institution with a large credit stake in the company, experience lower than average post-IPO share price development. They discover that when investment funds affiliated to the underwriting financial institution purchase the underwritten shares extensively, share price development is even more negative, suggesting the existence of conflicts of interest.

It seems that research focusing on directly finding economies of scope or conflicts of interest is also mixed. However, based on the literature it can be concluded that evidence on the existence of either one is limited.

2.3. Credit Default Swaps

Credit default swaps (CDSs) provide insurance against the default risk of the referred financial institution. The buyer of the CDS insurance pays the seller periodic payments until a credit risk event occurs or until the maturity of the CDS contract. The annual payment, referred as the CDS spread, is expressed as basis points of the value of the CDS contract. If a credit event occurs (i.e. the default of the underlying financial institution) the buyer of the CDS is compensated for the losses, equalling the difference between the par value of the underlying bond and its market value after the credit event. The CDS spread provides a robust measure of the default probability (credit risk) of the referred financial institution.

According to Jorion and Zhang (2007) using CDS spreads instead of the spread between the corporate and Treasury bond yields is preferable to measure credit risk. They argue that bond measures are sensitive to the choice of the risk-free rate and other risk factors not related to credit risk. Chen, Lesmond and Wei (2007) find that bond yield spreads are strongly related to liquidity measures, like bond bid-ask spreads. Longstaff, Mithal and Neis (2005) find that corporate yield spreads have a credit risk and liquidity component. Moreover, Blanco, Brennan and Marsh (2005) find that CDSs have a clear lead over bond yield spread in the price discovery process. They also suggest that CDSs are cleaner indicators over bonds yield spreads on credit risk, concluding that CDSs are more useful in analysing credit risk.

Equity based risk measures contain some credit risk information (Vassalou, Xing 2004). However, when measuring credit risk CDSs are preferable over equity market measures, since credit events can imply different results in CDS and equity markets (Jorion, Zhang 2007). As an example, increases in leverage implies increases in credit risk and higher CDS spreads, however increases in leverage can transform value from bond holders to equity holders implying increases in equity value (Jorion, Zhang 2007). Furthermore, equity risk measures capture multiple risk factors instead of only credit risk (Fama, French 1993). Thus using CDS spread over equity measures is preferable to measure credit risk.

3. Data and Methodology

3.1. Selected Financial Institutions

In this study, the data sample consists of 51 financial institutions from Europe and the US, 42 from Europe and 9 from the US. The European financial institutions are selected from all of the listed financial institutions from the EU 15 countries, Norway and Switzerland. From the listed institutions, only the financial institutions with traded credit default swaps are included in the sample. From the US only financial institutions with liquid credit default swaps are included in the sample⁵. The time span of the data sample ranges from 2007 to 2014, because most of the currently traded CDSs started trading during 2007 when new CDS restructuring clauses were implemented. The data is in panel data form with 51 financial institutions during 2007-2014. However, the data is not fully balanced as 14 financial institution year observations are excluded due to missing data points, reducing the financial institution year observations from 408 to 394.

3.2. Credit Default Swap Spreads

3.2.1. Financial Institution level Credit Default Swap Spreads

The credit default swap spreads used in this study have been collected from Datastream Advanced database. For each financial institution daily CDS spreads, expressed in basis points, are collected from 2007 to 2014. To measure the yearly credit risk of the financial institutions as the dependable variable, the average of the daily senior debt CDS spreads within each year are used. CDS spread can be obtained in different maturities, in this study the most liquid⁶ 5 year maturity is used. However, for the sake of robustness also 1 and 10 year maturities are used. CDS spreads can also be obtained for the junior debt of the financial institutions. Junior debt is considered as mezzanine capital and can have equity risk characteristics in addition to credit risk. Because of this, only senior CDS spreads are used.

⁵ The liquid US financial institutions have been selected based on reports from the Kamakura Corporation and iBanknet.

⁶ According to the Bank of Finland “*The determinants of global bank credit-default-swap spreads*” the 5 year CDS maturity is the most liquid maturity.

Different contractual restructuring clauses have been introduced to CDSs, which define the credit events that trigger the CDS settlement. Overall, there are multiple credit events that trigger payments from the seller of the CDS contract to the buyer; bankruptcy, failure to pay, repudiation/moratorium, obligation acceleration, obligation default and restructuring. The restructuring credit event is the most difficult credit event to contract. Major differences between different CDS contract clauses arise from the contractual handling of the restructuring credit event. The full-restructuring (CR) clause used to be the standard clause, under which any debt restructuring event is considered as a credit event. Under the CR clause, debt restructurings that did not cause losses to the bond holders could constitute as credit events and trigger payments to the CDS protection byers. As a response to the issue, the modified restructuring clause (MR) was introduced in 2001, to limit credit events that did not cause losses. Under the MR clause restructuring agreements still constitute as credit events (excluding the restructuring of bilateral loans), however the MR clause limited the obligations deliverable in restructuring agreement credit events to maturity under 30 months. In 2003 the modified-modified (MM) restructuring clause was introduced, which increased the maturity limit of the deliverable obligations to 60 months. In this study the MM restructuring clause is primarily used, which is common in Europe and was found for most of the financial institutions used in this study.⁷ However, for 3 European financial institutions the CR clause is used and for 2 US financial institutions the MR restructuring clause is used⁸. It is not believed that the use of non-MM clauses for 5 financial institutions impacts the results. The spread difference between clauses is not economically different, on average CR clause is priced 3.4 basis points higher than MR and MM clause trades between CR and MR.⁹

Figure 1 describes the yearly CDS spread observation for the 5-year maturity senior debt CDS spreads used in the sample. Each yearly data point represents the yearly average of the daily CDS spreads of a single financial institution within a specific year.

⁷ Information on credit default swap contractual terms and clauses has been obtained from the Bank of International Settlements, Thompson Reuters and Markit.

⁸ The 5 financial institutions do not have MM clause CDSs

⁹ Bank of International Settlements was used as the source for the spread differences between restructuring clauses.

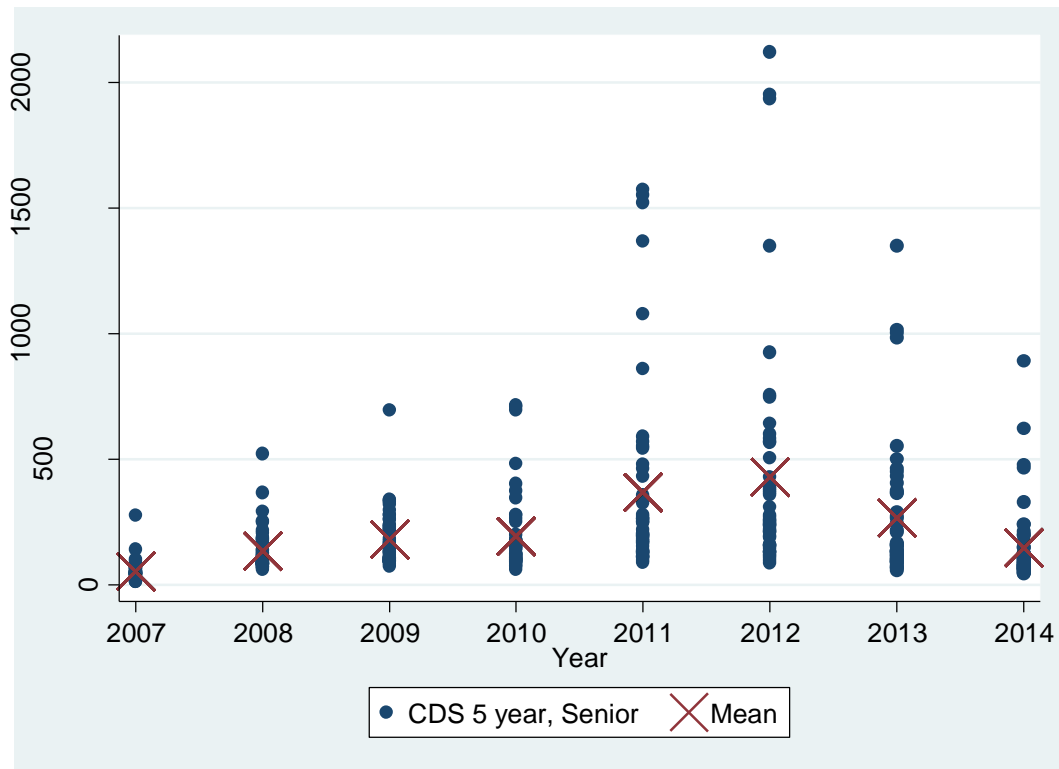
Figure 1 – CDS Spreads

Figure 1 shows that the mean of the yearly CDS spreads has increased from 2007 to 2012 and that the standard deviation has also increased during the sample period. It seems that some financial institutions have experienced large increases in credit risk, increasing the cross sectional variation at the financial institution level. After 2012, the mean CDS spread and the cross sectional variation have decreased. The observations are consistent with the evolution of the financial crisis which started in 2008 and forced financial institutions to write-off assets. Furthermore, the Eurozone debt crisis started in 2009 and reached its peak during the 2011-2012, threatening the stability of the European financial sector. As most of the financial institutions in the sample are EU based, a clear spike in the CDS spreads can be seen in 2011-2012.

3.2.2. Credit Default Swap Banking Sector Index Spreads

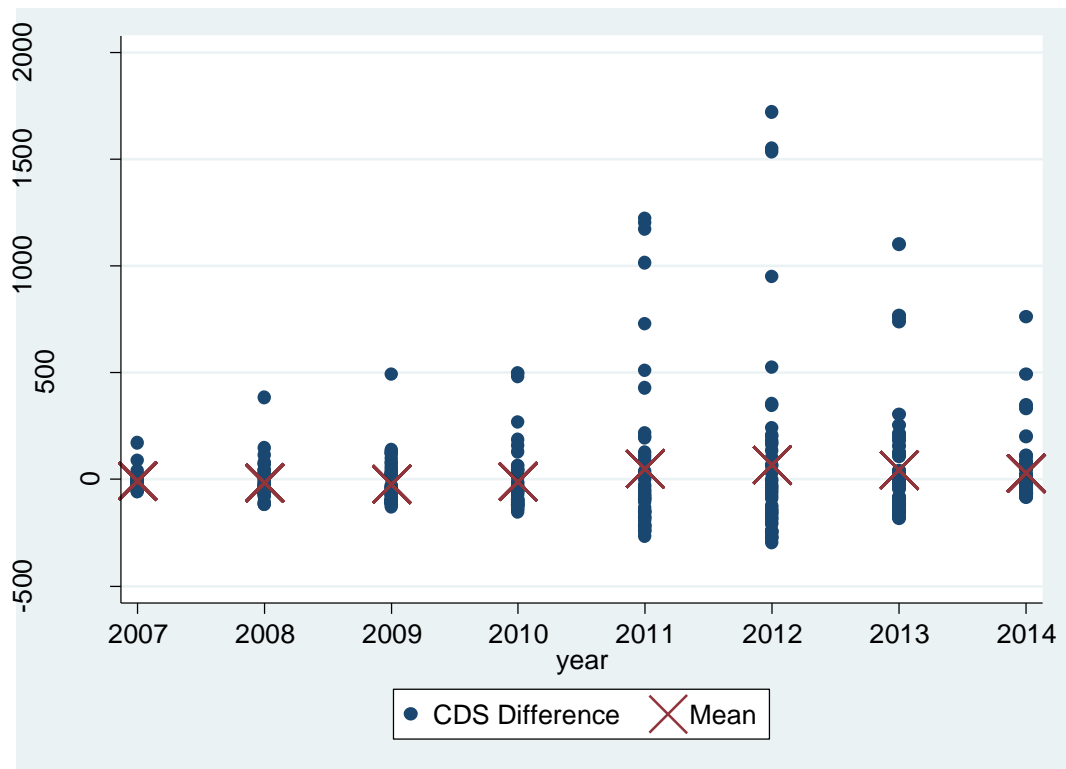
Daily banking sectors CDS index spreads are collected from Datastream Advanced database between 2007 and 2014. For each year the yearly averages of the daily CDS index spreads are calculated, to represent the average credit risk in the banking sector during the year. The index acts as a benchmark to which the financial institution level CDS spreads are

compared. The motivation is to also use the difference between the financial institution level CDS spread and the banking index CDS spread in the regressions as the dependable variable. For European financial institutions the *ds europe banks 5y cds index* is used and for US financial institutions the *ds na banks 5y cds index* is used as the benchmark. As both CDS indexes have 5-year maturities also the financial institution level 5-year CDS spreads are used in the difference. In order to calculate the difference the following method is used.

$$CDS\ DIFF_{.nt} = CDS_{nt} - CDS\ index_t$$

If the CDS spread of the financial institution is lower (higher) than the banking sector CDS index spread, the difference is negative (positive). Meaning that negative (positive) CDS difference values for a financial institution result in lower (higher) than the industry's average credit risk. Figure 2 demonstrates the annual financial institution level CDS spread differences.

Figure 2 – CDS Difference



3.3. Functional Focus and Diversification Measures

The explanatory and control variables used in this study are derived from the income statement and balance sheet data of the financial institutions in the sample, the data has been

collected from the annual financial reports of the 51 financial institutions in the sample¹⁰. The data ranges from 2007 to 2014 and includes only annual data.

The functional focus and diversification measurements used in this study follow methods used in previous literature. Income and asset based methods are used to measure both functional focus and diversification. Income based method of measuring functional focus measures which income stream of the financial institution is larger, commercial banking or investment banking income stream. Income based functional diversification measures how much the income streams from commercial banking and investment banking are diversified. The asset based functional focus method measures which commercial banking or investment banking assets are larger and diversification how well the two asset categories are diversified in the balance sheet.

3.3.1. Income Based Measures for Functional Focus and Diversification

Measurements of income diversification and focus compare the proportion of net interest income with the non-interest income of the financial institution. Net interest income is mostly generated from the traditional commercial banking activities of the financial institution. Net interest income is calculated by deducting interest expense (mostly the interest financial institutions pay on the financing of their lending activities) from the interest income (interest financial institutions receive from their lending activities), thus net interest income mostly captures the extent of the financial institutions commercial banking activities. Non-interest income captures income generated by investment banking¹¹ activities of the financial institution, which includes net fee income, net commission income, net trading income, net insurance income and other income. Non-interest income is mostly generated from investment banking, insurance, underwriting and distributing securities, securitizing assets, mutual fund services and other financial services.

3.3.1.1. Net Interest Income Share

¹⁰ The annual statements have been obtained from the websites of the financial institutions and the data has been collected by hand from the statements.

¹¹ Even though referred to as investment banking, non-interest income can include income from other financial services (e.g. insurance). The focus of the financial institutions is divided between commercial banking and investment banking for the sake of clarity and since non-interest income mostly contains income from investment banking.

Net interest income share, used by (Baele, De Jonghe & Vander Venet 2007), simply measures the share of net interest income to total operating income. Higher the ratio more the financial institution generates its income from commercial banking. With a lower ratio, more income is generated from investment banking activities. Thus the net interest income share measures the functional focus of the financial institution.

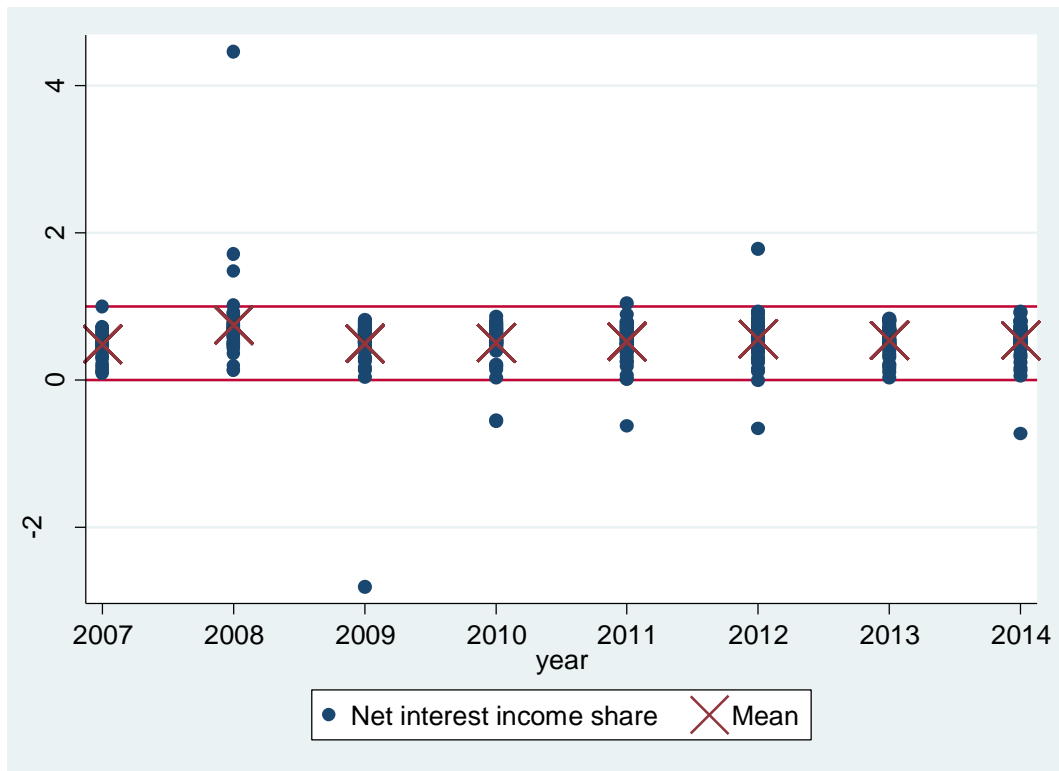
$$\text{Net interest Income Share} = \frac{INT}{TOR}$$

Where:

INT = net interest income

TOR = total operating income = net interest income + non-interest income

The net interest income share does not properly capture functional diversification. The method describes to which activity the financial institution is focusing on, not diversification. The high (low) values of the measurement imply that the financial institution is focused on commercial banking (investment banking). Only medium values imply that the financial institution would be diversified. However, since commercial banking and investment banking activities can have different risk profiles, the method is used to determine if there is a difference in credit riskiness between commercial banking and investment banking. Furthermore, when the functional diversification is examined, the net interest income share is used as a control variable.

Figure 3 – Net Interest Income Share

Judging by the means of the net interest income share ratio, the share of net interest income has slightly increased from 0.48 in 2007 to 0.53 in 2014. Moreover, the mean of the year 2008 (0.74) is clearly higher than the means of other years, this can be explained by the outbreak of the financial crisis. Based on the 2008 financial statements of the financial institutions in the sample, many financial institutions wrote off large amounts of trading assets and their net trading income was negative. This leads to small or even negative non-interest income, increasing the net interest income share ratio. As evident from Figure 3 not all observations lie between 0 and 1 as intended. The larger than 1 values result from the negative non-interest income, making net interest income larger than total operating income. The negative values result from negative total operating income. The larger than 1 and negative values are not removed from the sample in order to avoid bias.

3.3.1.2. Income Diversification

The income diversification measure used by (Laeven, Levine 2007) captures income diversification more effectively than the net interest income share. Income diversification is

calculated by taking the absolute value from, the difference between net interest income and non-interest income divided by total operating income.

$$INC.DIV = 1 - \left| \frac{INT - NON.INT}{TOR} \right|$$

Where:

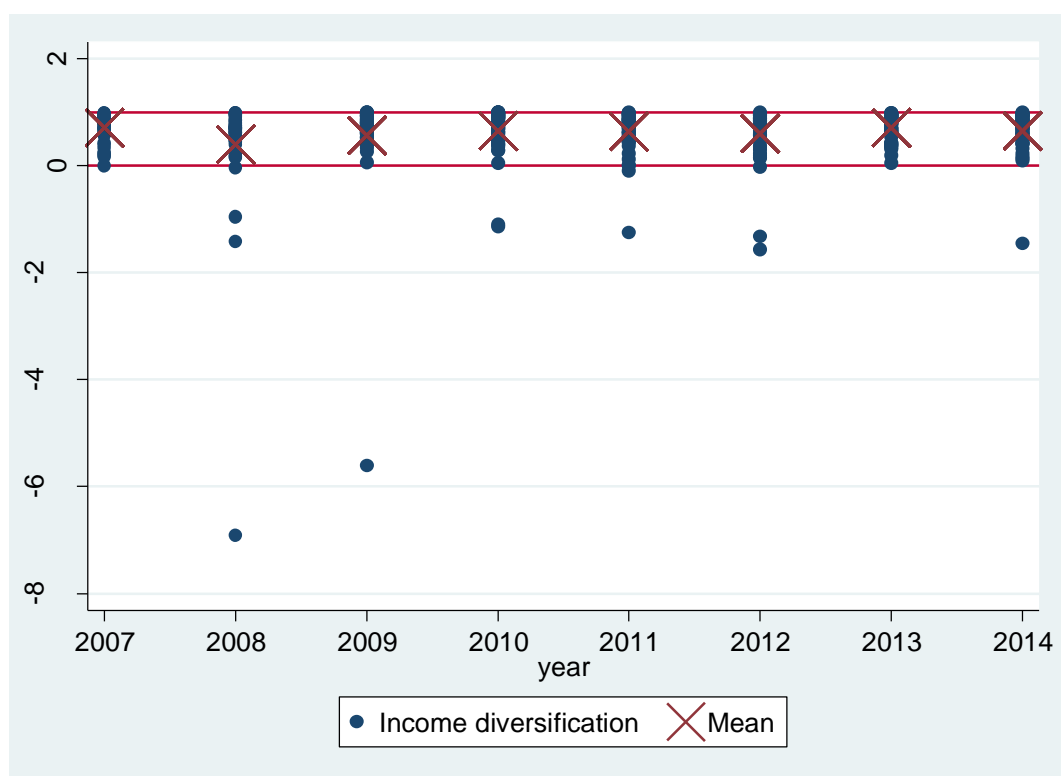
INC.DIV = income diversification

INT = net interest income

NON.INT = non-interest income

TOR = total operating income

The income diversification measurement takes values between 1 and 0, where 0 implies no diversification since income is generated exclusively from net interest income or non-interest income, 1 implies high diversification since income is generated evenly from both net interest income and non-interest income. The measure captures income diversification effectively for regression purposes, since the relationship between high and low diversification is linear. The measure returns the same values regardless which income stream is larger, thus only focusing on diversification. Income diversification is the main explanatory variable used in this study to measure the functional diversification.

Figure 4 – Income Diversification

The mean of the income diversification measure is the highest during 2007 at 0.71. During 2008, the mean income diversification decreased to 0.40, due to the decreases in non-interest income. The mean of income diversification varies between 0.57 and 0.71 during 2009-2014. Observations indicate a relatively high level of average diversification among the financial institutions in the sample. Figure 4 shows that the financial institutions in the sample have wide cross sectional variation in functional diversification, as the income diversification measures are quite evenly distributed between 0 and 1. Again negative observations result from negative non-interest income or total operating income.

3.3.1.3. Herfindahl-Hirschman Income Diversification

Second method used to measure functional diversification is the application of modified Herfindahl-Hirschman index to income diversification, also used by (Elsas, Hackethal & Holzhäuser 2010, Stroh, Rumble 2006a). The HH income diversification is used in the robustness section of this study.

$$HH\ INC.\ DIV = 1 - \left(\left(\frac{INT}{TOR} \right)^2 + \left(\frac{NON.INT}{TOR} \right)^2 \right)$$

Where:

HH INC.DIV = Herfindahl-Hirschman income diversification

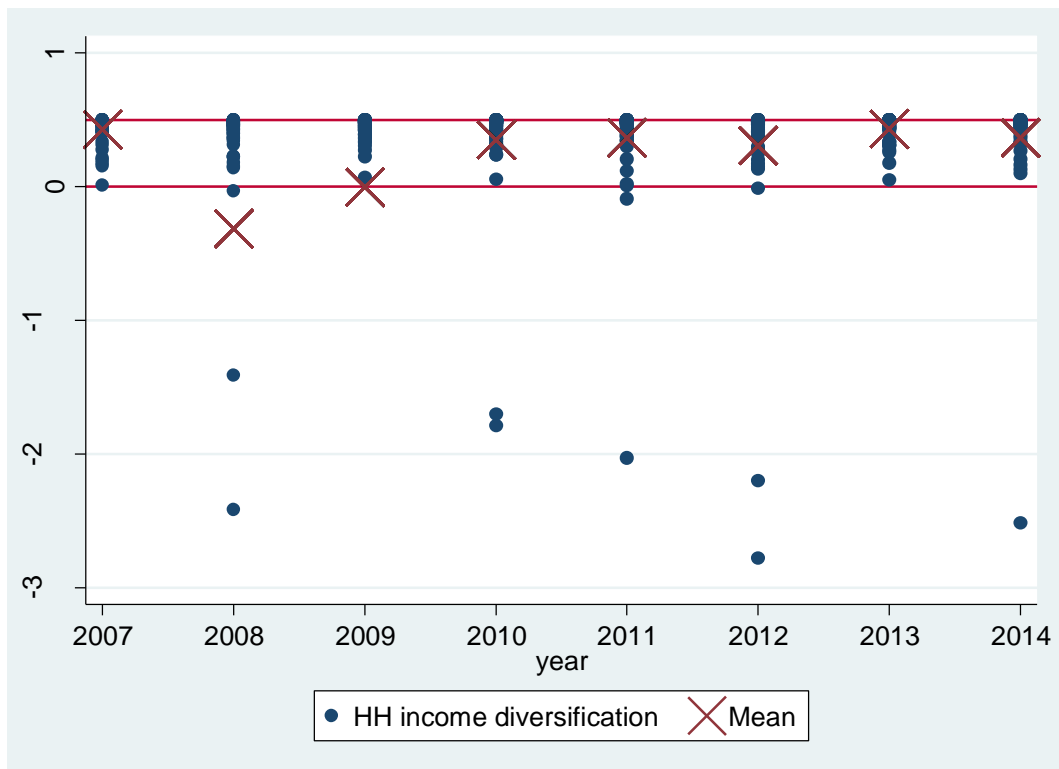
INT = net interest income

NON.INT = non-interest income

TOR = total operating income

The HH income diversification measure can take values between 0 and 0.5, 0 implying the financial institution is fully specializing in commercial banking or investment banking and 0.5 implying high diversification where income streams are evenly divided between the 2 income streams. The measure returns the same values regardless which income stream is larger, only focusing on diversification.

Figure 5 – HH Income Diversification



The HH income diversification measure tells the same story as the income diversification measure, diversification among the financial institutions in the sample seems quite high based on the means. The negative means of the 2008 and 2009 result from two highly negative observations for one financial institution -30.80 in 2008 and -21.35 in 2009. The extreme values have not been plotted in Figure 5, in order to increase readability. Negative non-interest income and total operating income result in negative HH income diversification measures.

3.3.2. Asset Based Measures for Functional Diversification

Similar methods are used to measure the functional focus and functional diversification based on the balance sheets of the financial institutions. The purpose of the asset measurements is to measure the share and mix of commercial banking assets and investment banking assets in the balance sheet of the financial institution. Since traditional commercial banking is mostly lending based, loans in the balance sheet are categorised as commercial banking assets. Other earning assets mostly include trading assets, derivatives, insurance assets, different investments and other financial assets and are categorized as investment banking assets. Non-earning assets such as cash, non-investment real-estate, equipment, intangible assets and tax assets are not included in neither category nor total earning assets.

3.3.2.1. Loan Share

Loan share measures the share of loans to total earning assets and is used to measure the functional focus of the financial institutions based on its balance sheet.

$$\text{Loan Share} = \frac{\text{LOAN}}{\text{TEA}}$$

Where:

LOAN = net loans to customers

TEA = total earning assets (loans to customers + other earning assets)

Similarly to the net interest income share, high (low) values imply high focus on commercial banking (investment banking) and mid values high diversification. The loan share is used to determine if the credit risk between commercial banking and investment banking is different and as a control variable.

Figure 6 – Loan Share



The loan share measure seems to be evenly distributed between 0.8 and 0.2, with mean values ranging between 0.54 and 0.57. Meaning that on average the earning asset base of financial institutions contains more loans than other earning assets. The mean loan share was 0.54 in 2007, in 2008 the value increased to 0.56 probably since financial institutions had to write off other earning assets, like trading assets, due to the financial crisis. Eventually in 2014 the mean share of loans to total earning assets decreased back to 0.54. Not a single financial institution's asset base consists solely of loans. Two financial institutions sole assets base contained only other earning assets during 2007 and 2009, with a slight increase of loans during 2009 and 2014. The financial institutions in question are US based investment banks Goldman Sachs and Morgan Stanley.

3.3.2.2. Asset Diversification

The asset diversification follows the same methodology as the income diversification measure and is used to measure functional diversification.

$$AS.DIV = 1 - \left| \frac{LOAN - OEA}{TEA} \right|$$

Where:

AS.DIV = asset diversification

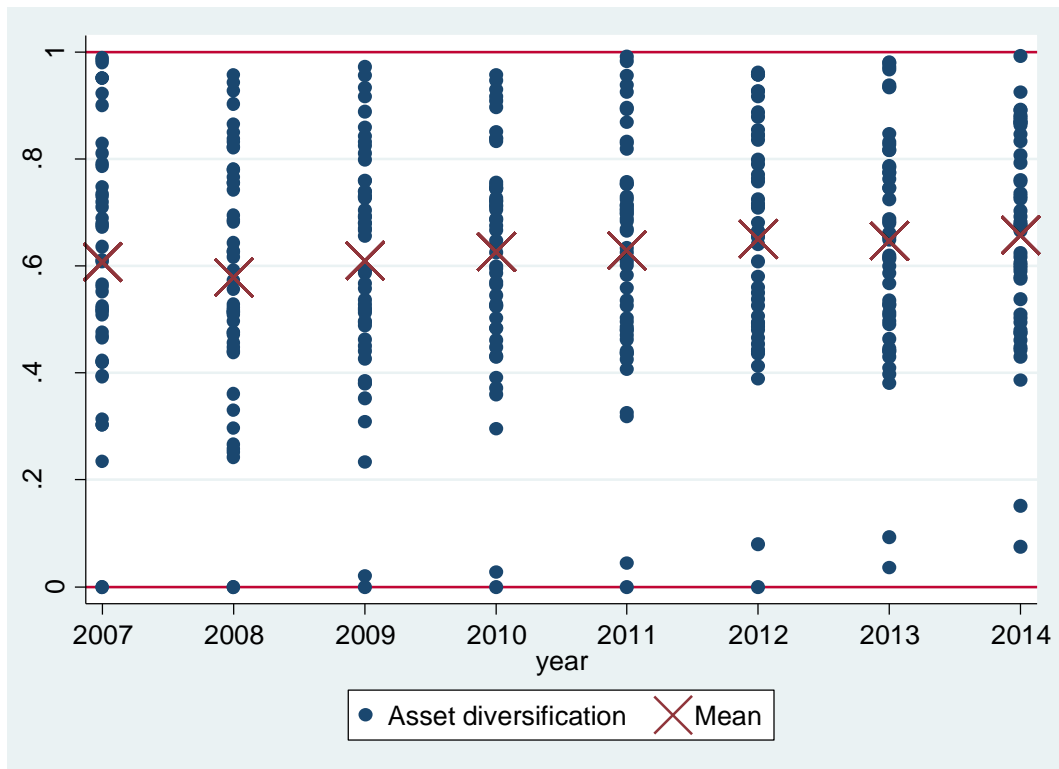
LOAN = net loans to customers

OEA = other earning assets

TEA = total earning assets (loans to customers + other earning assets)

The asset diversification measurement takes value between 1 and 0, where high (low) values imply high (low) functional diversification.

Figure 7 – Asset Diversification



The means of the asset diversification measure show a slight upward trend, the asset diversification has increased from 0.61 in 2007 to 0.66 in 2014. However, in 2008 asset diversification decreased, since the share of loans to total assets increased, decreasing diversification. Again the sample shows a wide array of different diversification observations.

3.3.2.3. Herfindahl-Hirschman Asset Diversification

The HH asset diversification measure uses the modified Herfindahl-Hirschman index method. The HH asset diversification is used in the robustness section of this study to measure functional diversification.

$$HH\ AS.\ DIV = 1 - \left(\left(\frac{LOAN}{TEA} \right)^2 + \left(\frac{OEA}{TEA} \right)^2 \right)$$

Where:

HH AS.DIV = Herfindahl-Hirschman asset diversification

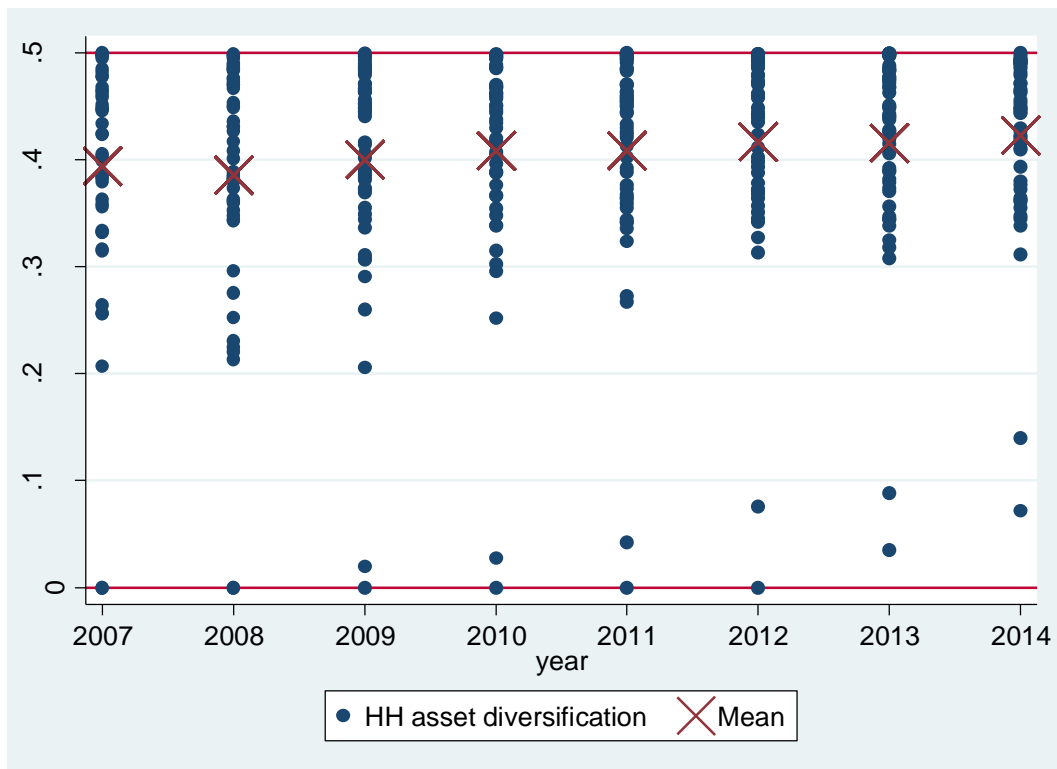
LOAN = net loans to customers

OEA = other earning assets

TEA = total earning assets (loans to customers + other earning assets)

The HH asset diversification measure can take values between 0 and 0.5, high (low) values implying high (low) diversification.

Figure 8 – HH Asset Diversification



3.4. Control Variables

In order to control for other potential explanatory factors for the credit risk of financial institutions, I employ several control variables in the panel data regressions. The used control variables follow previous literature¹². The control variables have been obtained from the annual financial statements of the financial institutions in the sample.

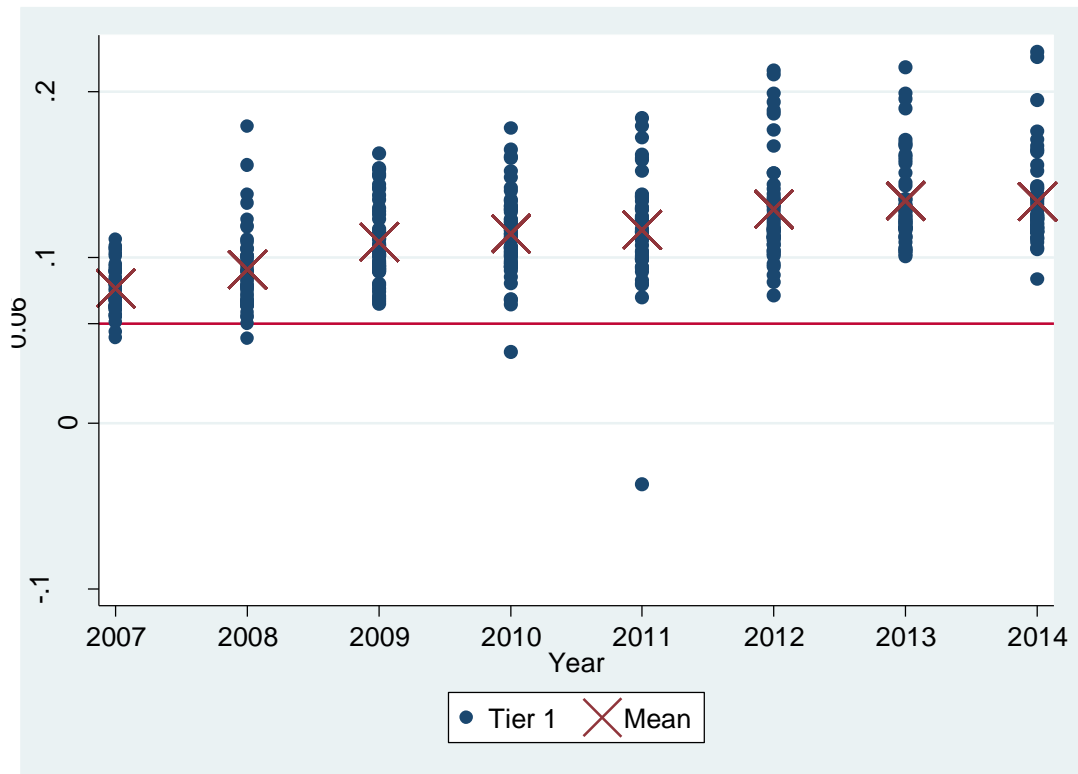
3.4.1. Tier 1 Capital Ratio

First major control variable is the tier 1 capital ratio of the financial institutions. The tier 1 capital ratio describes the financial strength and riskiness of the financial institution and is similar with the traditional leverage ratio. Tier 1 is used for regulatory purposes. The tier 1 capital ratio is calculated by dividing the core equity capital (or tier 1 capital) with the risk weighted assets of the financial institution. The risk weighted assets include all risky assets of the financial institution with different credit risk weights for each asset class. Central banks or national institutions in charge for regulating financial institutions develop the weights for different asset classes, which often follow the international Basel regulatory guidelines. Financial institutions with a tier 1 capital ratio under 6%¹³ are seen as undercapitalized and are forbidden to pay out any dividend or management fees. High tier 1 ratio implies lower riskiness and stronger capital structure of the financial institution. Thus the tier 1 ratio is expected to be negatively related to CDS spreads and credit risk.

¹² Some control variables follow methodology used in a study published by the Bank of Finland “The determinants of global bank credit-default-swap spreads”

¹³ Different limits exists in different regulatory jurisdictions, 6% is the most commonly known.

Figure 9 – Tier 1

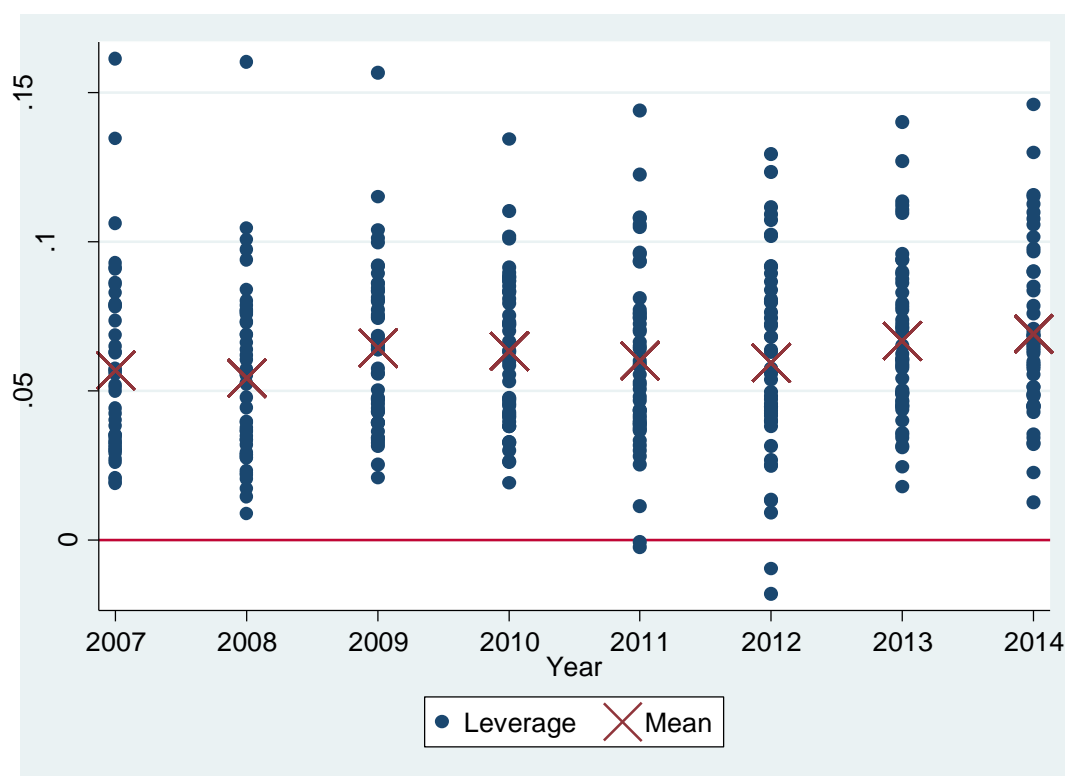


For the financial institutions in the sample, the mean tier 1 has steadily increased from 0.081 in 2007 to 0.134 in 2014. The increases in tier 1 capital ratio can be explained by the efforts of central banks to increase capital requirements, especially the Basel committee has been introducing stricter capital requirements for the largest global financial institutions. Only a small number of financial institutions in the sample have had a tier 1 ratio under 0.06 and only one negative ratio can be observed.

3.4.2. Leverage Ratio

In addition to the tier 1 ratio, the traditional leverage ratio is employed as a control variable. Leverage should capture the effect of the actual financial structure on the credit risk of financial institutions. The leverage ratio has been calculated by dividing total equity with the total assets of the financial institutions. In this case, a high leverage ratio implies that lower amount of debt is used in the financing of the financial institution. Thus a high leverage ratio results in stronger capital structure and lower risk for the financial institution. The relationship between the leverage ratio and credit risk is expected to be negative.

Figure 10 – Leverage Ratio



The mean leverage ratio has also increased from 0,057 in 2007 to 0,069 in 2014. Negative values in 2011 and 2012 are observed as the National Bank of Greece, Dexia and Eurobank Ergasias had negative equity values during those years. The means of the leverage observations increased by 21% between 2007 and 2014, as the tier 1 observations increased by 65%. Since the increase in the leverage ratio is much smaller, it can be interpreted as an indication that financial institutions have increased their tier 1 capital ratios by both increasing the relative amount of equity and by shifting their asset base towards less risky assets (decreasing risk weighted assets).

3.4.3. Cost to Income Ratio

The cost to income ratio is widely used to measure the profitability of financial institutions. The ratio is calculated by dividing the operational costs (salaries, SQ&A, the depreciation of non-financial assets and other costs) with the total net operating income (net interest income and non-interest income). Provisions for credit losses (losses and expected losses on loans) are not included in the ratio. Lower (higher) the ratio more (less) profitable the

financial institution is. If the ratio is one or higher, the operational costs of the financial institution exceed the income and the financial institution is unprofitable.

Figure 11 – Cost to Income Ratio



During 2008, the profitability of financial institutions was low, potentially because many financial institutions occurred losses in trading and wrote down financial assets. Negative values indicate that the total operating income of the financial institution has been negative.

3.4.4. Size

The effect of the size of financial institutions on risk and the stability of the banking sector (so called too big to fail doctrine) has been much debated. According to research, if large financial institutions are at risk of default, governments might save the financial institution in order to avoid the adverse impacts of defaulting financial institutions on the economy (O'Hara, Shaw 1990). The size of the financial institutions is used as a control variable, measured by the natural logarithm of the total assets of the financial institutions. The relationship between size and credit risk is expected to be negative, as larger financial institution should have lower probabilities of default as governments might save them from default.

3.4.5. Loan Loss Provisions to Net Loans

Financial institutions make credit loss provisions in the income statement, based on loan losses and expected losses. Through the provisions, they write off assets (mostly loans) based on expected defaults on customer loans. The ratio of credit loss provisions to net loans measures the credit risk exposure of the loan portfolio of financial institutions and is included as a control variable. A high ratio implies high credit risk in the loan portfolio. The loan loss provisions ratio is expected to be positively related with CDS spreads and credit risk.

3.4.6. Interest Expense to Liabilities

The interest expense divided by the total liabilities ratio of the financial institution measures the financing cost of the financial institution. Even though not a market based measure, it can indicate the credit risk of the financial institution. Increased credit risk should in theory increase the interest rate margin available for the financing of the financial institution, as financiers demand compensation on the increased riskiness. Thus, the relationship between the interest expense ratio and CDS spreads is expected to be positive.

3.4.7. Liquid Assets to Total Assets

A part of credit risk for financial institutions can result from liquidity risk. If the financial institution does not have enough liquid assets, it may not be able to pay its payables in adverse market conditions. In the case of a bank run, depositors and other financiers of the financial institutions might withdraw larger than expected amounts of capital from the financial institution. If the financial institution does not have enough liquid assets to fulfil the unexpected withdrawals, it might face bankruptcy even if its operations were profitable. Thus in order to control for liquidity risk the liquid assets to total assets ratio is used as a control variable. Only the most liquid assets are included in the ratio, cash and deposits in central banks. A higher ratio should indicate lower liquidity risk. A negative relationship between the liquidity ratio and CDS spreads is expected.

3.5. Summary Statistics

Table I – Summary Statistics 2007-2014

Variable	N	Std. Dev.	Mean	Min	P25	P50	P75	Max
CDS 5y	397	280.39	223.49	14.79	87.08	129.40	243.77	2121.58
CDS 1y	393	349.02	186.71	3.16	39.33	83.61	173.27	3024.38
CDS 10y	393	235.05	217.52	4.77	96.81	145.37	247.44	1824.28
CDS difference	397	250.70	16.92	-294.17	-96.10	-36.73	28.84	1721.43
Income share	397	0.359	0.548	-2.81	0.47	0.57	0.68	4.46
Income diversification	397	0.615	0.618	-6.91	0.53	0.73	0.89	1
HH Income div	397	1.94	0.238	-30.79	0.39	0.46	0.49	0.5
Loan share	397	0.209	0.556	0	0.41	0.62	0.72	0.884
Asset diversification	397	0.216	0.626	0	0.49	0.65	0.78	0.993
HH Asset div.	397	0.101	0.407	0	0.37	0.44	0.48	0.5
Tier 1	394	0.032	0.114	-0.037	0.09	0.11	0.13	0.224
Leverage	397	0.029	0.062	-0.018	0.04	0.06	0.08	0.161
Cost/income	397	0.525	0.651	-3.47	0.53	0.62	0.72	6.80
Ln(size)	397	1.18	12.78	10.36	11.73	12.94	13.92	14.92
Provisions/loans	397	0.014	0.012	-0.004	0.003	0.009	0.016	0.095
Interest expense /liabilities	397	0.017	0.022	0.003	0.012	0.017	0.029	0.163
Liquid assets /assets	397	0.030	0.034	0.000	0.012	0.025	0.051	0.163

4. Hypothesis

Previous literature on the effect of functional diversification on the riskiness of financial institutions has provided conflicting results. The results vary based on geographical scope, the methods of measuring diversification and methods on measuring market valuations and risk. Some studies find that diversification increases (decreases) riskiness (market values) of financial institutions and other studies the opposite. These studies use equity based risk and value measurements to determine the riskiness or value of the financial institutions and the relationship between functional diversification. To determine whether financial institutions

should be allowed to functionally diversify (combine commercial banking with investment banking) the most relevant risk factor from the regulatory perspective is credit risk. Using the equity based risk and valuation methods multiple risk factors are captured, instead of only credit risk. Regulation is mostly interested in to protect the economy from the negative impacts of banking failures and defaults. This is why this study concentrates on measuring financial institution level risk, using CDS spreads that directly capture the credit risk of financial institutions.

Before determining whether functional diversification impacts the credit risk of financial institutions, the impact of functional focus (commercial or investment banking) is examined. Since diversification benefits can arise from lower exposure to a riskier function, not diversification itself. Moreover, when examining functional diversification, the functional focus of the financial institution is controlled for.

4.1. Research Question 1 – Are Financial Institutions Focusing More on Investment Banking More Risky Than Those Focusing on Commercial Banking

Previous research has found that financial institutions focusing on fee generating investment banking are riskier than those focusing on net interest income generating commercial banking activities. No studies have used bond or CDS data to measure credit risk. However, in light of previous research it is hypothesised that a lower share of net interest income or loan share leads to higher credit risk, resulting in higher (lower) credit risk for financial institutions focusing more on investment banking (commercial banking).

H1 – Financial institutions focusing on investment banking have higher credit risk.

In order to answer the first research question, the following OLS panel data regressions are developed. At first random effects OLS panel data regressions will be used with the following notation, to answer if financial institutions focusing more on investment banking have higher credit risk.

$$CDS_{nt} = \beta_0 + \beta_1 \text{Interest share} + \beta_2 \text{Loan share}_{nt} + \beta_3 \text{Tier1}_{nt} + \beta_4 \text{LEV}_{nt} + \beta_5 \frac{\text{Cost}}{\text{INC}_{nt}} + \beta_6 \text{LN}(\text{Size})_{nt} + \beta_7 \frac{\text{PROV}}{\text{Loans}_{nt}} + \beta_8 \frac{\text{INT.EXP}}{\text{Liabilities}_{nt}} + \beta_9 \frac{\text{Liquid AS.}}{\text{Assets}_{nt}}$$

The dependent variable is the credit risk measure of financial institution n at year t , the preferred credit risk measure is the yearly average of the daily CDS 5-year maturity spread observations. For the sake of robustness, also 1 and 10-year maturity spreads and the difference between the financial institution level 5-year CDS spread and the banking sector CDS index are used to measure credit risk. The explanatory variables net interest income share and loan share describe whether the financial institution focuses more on commercial banking or investment banking. For both explanatory variables, a negative coefficient is expected which would mean that financial institution focusing more on investment banking would have higher credit risk. In all of the regressions a set of previously defined control variables are used to control for potential other explanations for the cross sectional heterogeneity in the credit risk of financial institutions.

In the robustness section of his study, fixed effect panel data regressions are used. The unobserved and uncontrolled for characteristics of individual financial institution could impact the results when random effects panel data regressions are used. In fixed effect regressions, dummy variables for each financial institution (omitting the first financial institution in order to avoid the dummy variable trap) are developed in order to control for unobserved heterogeneity. The results gained from the random effect regressions could be driven by time effects, such as the financial crisis which took place during the sample of this study. Also year dummies are developed for the fixed effect regressions to control for time effects.

$$\begin{aligned}
CDS_{nt} = & \beta_0 + \beta_1 \mathbf{Interest\ share} + \beta_2 \mathbf{Loan\ share}_{nt} + \beta_3 \mathbf{Tier1}_{nt} + \\
& \beta_4 \mathbf{LEV}_{nt} + \beta_5 \frac{\mathbf{Cost}}{\mathbf{INC}_{nt}} + \beta_6 \mathbf{LN(Size)}_{nt} + \beta_7 \frac{\mathbf{PROV}}{\mathbf{Loans}_{nt}} + \beta_8 \frac{\mathbf{INT.EXP}}{\mathbf{Liabilities}_{nt}} + \\
& \beta_9 \frac{\mathbf{Liquid\ AS.}}{\mathbf{Assets}_{nt}} + \gamma_2 \mathbf{Financial\ institution}_2 + \dots + \\
& \gamma_{51} \mathbf{Financial\ institution}_{51} + \delta_2 \mathbf{Year}_2 + \dots + \delta_8 \mathbf{Year}_8
\end{aligned}$$

4.2. Research Question 2 – Are More Diversified Financial Institutions Less Risky?

As the previous literature is mixed based on the effect of functional diversification on the riskiness of financial institutions, this study takes the hypothesis that increased diversification lowers the credit risk of financial institutions. By implementing the portfolio theory (Markowitz

1952), different income sources which are not perfectly correlated, should decrease the default probability of financial institutions.

H2 – Financial institutions with higher functional diversification have lower credit risk.

In order to answer if more functionally diversified financial institutions have lower credit risk, income and asset diversification measures are used as the explanatory variables. For both variables negative coefficients are expected since diversification is expected to decrease credit risk. When the income and asset diversification measures are used in the regressions, the net interest income and loan share ratios are used as control variables. The net interest income and loan shares measures do not measure diversification, but the extent on which the financial institution focusses on commercial banking or investment banking. The income and asset diversification measures capture the level of diversification, since both methods give the same values regardless which commercial banking or investment banking the financial institutions focusses on. However, the credit risk levels of financial institutions focusing on commercial banking or investment banking can differ. Thus when using the diversification measures as explanatory variables, the net interest income and loan shares are used to control for the potential differences in credit risk levels based on the functional focus of the institution with the following notation.

$$\begin{aligned}
 CDS_{nt} = & \beta_0 + \beta_1 INC.DIV_{nt} + \beta_2 AS.DIV_{nt} + \beta_{10} Interest\ share_{nt} + \\
 & \beta_{11} Loan\ share_{nt} + \beta_3 Tier1_{nt} + \beta_4 LEV_{nt} + \beta_5 \frac{Cost}{INC}_{nt} + \beta_6 LN(Size)_{nt} + \\
 & \beta_7 \frac{PROV}{Loans_{nt}} + \beta_8 \frac{INT.EXP}{Liabilities_{nt}} + \beta_9 \frac{Liquid\ AS.}{Assets}_{nt}
 \end{aligned}$$

Similarly to the first research question, fixed effects regressions are also used for the second research question to increase robustness. The CDS spread difference, calculated as the difference between the financial institution level CDS spread and the banking sector CDS sector index, is used as the dependable variable for credit risk. The motivation is to further increase the robustness of the results. By using the CDS difference as the dependable credit risk measure, the regressions do not take into account the changes in the overall credit risk environment of the banking sector. When, only using the financial institution level CDS spreads, the changes in the overall credit risk environment are taken into account and may impact results. When the difference is used, bank level changes in credit risk are more evident and should improve results.

The following notation is used for research question 2, moreover the same CDS difference method is also used for the first research question.

$$\begin{aligned}
CDS_{nt} - CDS\ index_t = & \beta_0 + \beta_1 INC.DIV_{nt} + \beta_2 AS.DIV_{nt} + \\
& \beta_{10} Interest\ share_{nt} + \beta_{11} Loan\ share_{nt} + \beta_3 Tier1_{nt} + \beta_4 LEV_{nt} + \\
& \beta_5 \frac{Cost}{INC}_{nt} + \beta_6 LN(Size)_{nt} + \beta_7 \frac{PROV}{Loans}_{nt} + \beta_8 \frac{INT.EXP}{Liabilities}_{nt} + \beta_9 \frac{Liquid\ AS.}{Assets}_{nt}
\end{aligned}$$

5. Results

The results of this study suggest that the credit risk of financial institutions focusing more on investment banking does not differ from those focusing more on commercial banking. Financial institutions with higher income diversification have lower CDS spreads, even when the functional focus of the financial institution is controlled for. Functional diversification reduces the credit risk of financial institutions and the reduction in credit risk is not created by functional focus. I find no evidence that financial institution with less loans in their balance sheet have higher credit risk or that asset diversification impacts the credit risk of financial institutions. Thus only hypothesis 2 is accepted, but only with the income based measures for the functional diversification of financial institution.

5.1. Is Investment Banking More Risky Than Commercial Banking?

At first it is studied whether financial institutions with higher focus on investment banking have higher credit risk than those focusing on traditional commercial banking. Table II contains the regression results using the net interest income to total income and loans to total earning assets as explanatory variables to measure the functional focus of the financial institutions. The dependable variable is the financial institution level CDS spreads as a measure for credit risk.

Table II – Net Interest Income Share and Loan Share, Random Effects Regressions

$Net\ interest\ income\ share = \frac{INT}{TOR}$ and $Loan\ share = \frac{LOAN}{TEA}$. Mainly 5 year maturity CDS spread is used in the regressions as a dependable variable, for the sake of robustness also 1 and 10 year maturities are used. CDS spreads are yearly averages of daily financial institution level CDS spread observations. In the table columns represent individual regressions, with the coefficient of the corresponding variable presented first in the variable cell. Below the coefficient the corresponding t-value is presented in brackets.

	CDS 5y	CDS 5y	CDS 5y	CDS 1y	CDS 10y
Net interest income share	-5.85 (-0.13)	-4.31 (-0.10)		12.43 (0.2)	-24.4 (-0.59)
Loan share	10.45 (0.09)		5.31 (0.05)	-61.58 (-0.4)	17.97 (0.17)
Tier 1	1600.39 (3.79)***	1592.33 (3.89)***	1591.80 (3.81)***	1154.70 (2.03)**	1537.24 (4.07)***
Leverage	-4091.24 (-7.70)***	-4078.08 (-7.80)***	-4073.15 (7.71)***	-3925.70 (-5.34)***	-2833.91 (5.73)***
Cost/income.	31.31 (0.99)	30.33 (1.01)	28.21 (1.38)	48.13 (1.14)	38.44 (1.37)
LN(Size)	-108.54 (-5.69)***	-109.52 (-7.24)***	-108.64 (-5.76)***	-122.87 (-4.78)***	-88.44 (-5.09)***
Provisions/loans	7795.04 (9.00)***	7805.79 (9.04)***	7817.19 (9.05)***	9460.24 (8.03)***	6091.73 (7.78)***
Int. exp./liabilities	-1491.76 (-1.75)*	-1488.38 (1.75)*	-1489.82 (-1.75)*	-1002.32 (-0.88)	-1513.22 (-1.98)**
Liquid/assets	-407.03 (-0.89)	-400.56 (-0.9)	-403.35 (-0.89)	-402.07 (-0.66)	-356.49 (-0.87)
Constant	1609.89 (5.34)***	1627.94 (7.64)***	1612.23 (5.43)***	1789.06 (4.41)***	1303.24 (4.75)***
N	394	394	394	390	390
Prob>Chi ²	0.00	0.00	0.00	0.00	0.00
R ²	0.38	0.38	0.38	0.31	0.31

* $p < 0.1$; ** $p < 0.05$; *** $p < 0.01$

Both net interest income and the loan share ratios are not significant and the null hypothesis cannot be rejected. It seems that higher share of net interest income or loans does not reduce the CDS spreads of financial institutions, when the linear specification is used. Financial institutions that focus on offering commercial banking services are as risky as financial institutions focusing on investment banking in credit risk terms. The coefficient of both ratios change depending on which maturity of the CDS spreads is used, when using the 5

and 10-year maturities the coefficient for the net interest income share (loan share) ratio is negative (positive) and when using the 1-year maturity the sign changes for both ratios.

Based on the Table II results, some control variables have coefficient signs than were not expected and some as expected. The coefficient of the tier 1 ratio is positive and significant with a p value lower than 0.01 in all of the regressions using different CDS maturities, with the exception of the 1 year maturity where the coefficient is significant at 0.05 level. The positive coefficient ranges between 1155 and 1600, meaning that financial institutions with higher tier 1 ratios (indicating lower risk) have higher CDS spreads and higher credit risk. Using the 1600 coefficient, an increase of tier 1 ratio from the sample's first quartile tier 1 ratio of 0.094 to the third quartile 0.130 observation, translates to an increase of 57.60 basis points in the 5-year CDS spread. Compared with the mean value of the 5-year CDS spreads of 223.49 the change is quite significant. However, the relationship between the tier 1 ratio and CDS spreads is counterintuitive, increase in tier 1 should decrease risk and reduce CDS spreads. The result can be explained by the financial crisis, from 2007 to 2012 CDS spreads in the sample increased as the banking sector experienced a turbulent environment. In response to the banking crisis, regulators demanded financial institutions to increase their capital buffers and increase tier 1 ratios, as a result tier 1 ratios in the sample increased from 2007 to 2014. It is speculated that the positive coefficient of the tier 1 ratio is driven by the pressure from the regulators. Further examination is called for.

The cost to income measure of the profitability of financial institutions has a positive coefficient, meaning that a higher cost to income ratio, i.e. lower profitability, translates to higher CDS spreads as expected. However, as the coefficients in the regressions are not significant and the null hypothesis cannot be rejected.

The leverage ratio has negative coefficients ranging between -4091 and -3926 for the 5 and 1-year CDS maturities and -2834 for the 10-year maturity. All of the coefficients are significant with p values below 0.01. The signs of the coefficients are negative as expected, as a higher leverage ratio, i.e. more equity financing, reduces the CDS spreads and credit risk of financial institutions.

The natural logarithm of the size of the financial institution is negatively related to the CDS spread, larger financial institutions have lower CDS spreads and lower probability of default. All of the coefficients are significant with p values lower than 0.01. The results provide evidence supporting the too big to fail doctrine, as it seems that investors in the CDS market

price in the possibility that larger financial institutions are bailed out by the government or regulators in the face of potential default. However, no strong conclusions can be made as the too big to fail doctrine is not in the main focus of this study.

The loan loss provisions to net loans ratio has a positive coefficient as expected, a higher share of loan losses realized in the income statement translates to higher CDS spreads and credit risk for the financial institution. The results are significant with p values lower than 0.01.

The share of interest expense to total liabilities has negative coefficients, indicating that financial institutions with higher relative interest expense have counterintuitively lower credit risk. The results are not strongly significant as the p values for the 5-year maturity are significant only at the p value level of below 0.10, not significant for 1-year maturity and 0.05 significant in 10-year maturity. As in the case of the tier 1 ratio, the relationship can be potentially explained by the financial crisis. During the sample period of 2007-2014, the mean interest expense to loans ratio has decreased for the financial institutions in the sample and CDS spreads have increased. The decrease in interest expense can be explained by the decreasing market interest rates and access to cheap central bank financing by the financial institutions.

The share of liquid assets to total assets has a negative coefficient as expected. A higher share of liquid assets reduces credit risk as liquid assets protect financial institutions from negative shocks. However, the results are not significant and the null hypothesis cannot be rejected.

Table III presents results using the quadric and linear variables of the net interest income share and loan share ratios. The purpose is to examine if the relationship between credit risk and functional focus is non-linear.

**Table III – Net Interest Income Share and Loan Share with Quadric Terms,
Random Effects Regressions**

Net interest income share = $\frac{INT}{TOR}$ and *Loan share* = $\frac{LOAN}{TEA}$. Mainly 5 year maturity CDS spread is used in the regressions as a dependable variable, for the sake of robustness also 1 and 10 year maturities are used. CDS spreads are yearly averages of daily financial institution level CDS spread observations.

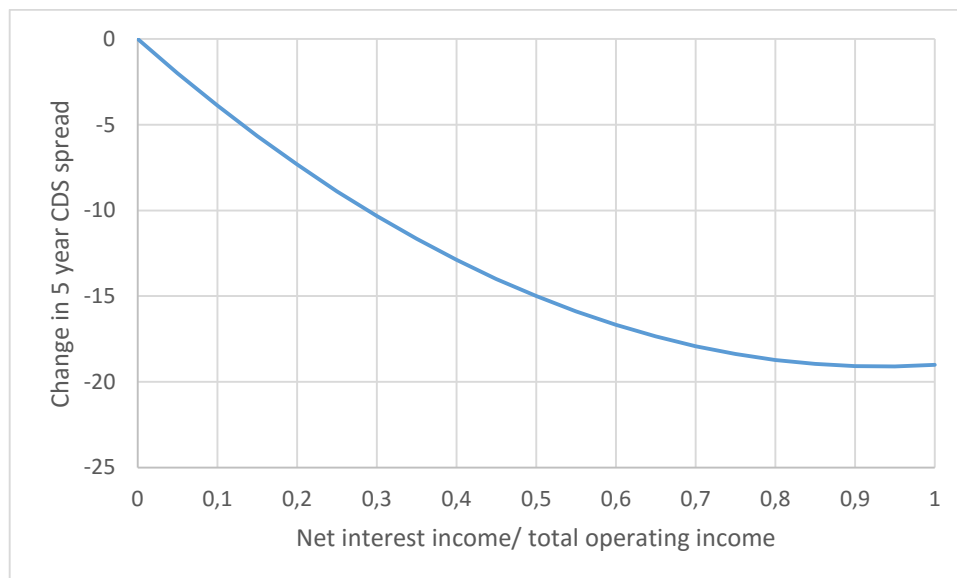
	CDS 5y	CDS 5y	CDS 5y	CDS 1y	CDS 10y
Net interest income share	-41.32 (-0.79)	-45.20 (-0.94)		-63.09 (-0.91)	-69.40 (-1.50)
(Net interest income share)²	22.28 (1.84)*	22.70 (1.90)*		41.65 (2.59)***	23.02 (2.15)**
Loan share	-180.83 (-0.56)		-197.84 (-0.64)	-272.00 (-0.62)	70.49 (0.24)
(Loan share)²	245.64 (0.72)		232.13 (0.69)	300.94 (0.64)	-23.74 (-0.08)
Tier 1	1659.42 (3.94)***	1636.23 (4.00)***	1593.63 (3.82)***	1293.41 (2.28)**	1590.64 (4.21)***
Leverage	-3941.17 (-7.57)***	-4044.18 (-7.69)***	-3972.66 (-7.66)***	-3859.49 (-5.27)***	-2727.94 (-5.56)***
Cost/income.	33.65 (1.05)	36.15 (1.20)	26.80 (1.30)	58.64 (1.37)	46.34 (1.63)
LN(Size)	-96.00 (-4.94)***	-108.14 (-7.08)***	-101.87 (-5.27)***	-107.14 (-3.95)***	-83.19 (-4.56)***
Provisions/loans	8047.90 (9.21)***	7711.08 (8.95)***	8123.20 (9.30)***	9508.35 (8.03)***	6085.76 (7.71)***
Int. exp./liabilities	-1413.03 (-1.69)*	-1597.89 (-1.87)*	-1311.63 (1.57)	-1109.45 (-0.98)	-1520.96 (-2.01)**
Liquid/assets	-500.42 (-1.11)	-401.89 (-0.90)	-475.03 (-1.06)	-494.81 (-0.82)	-415.19 (-1.03)
Constant	1472.77 (5.01)***	1615.20 (7.52)***	1546.25 (5.37)***	1600.99 (3.95)**	1212.41 (4.45)***
N	394	394	394	390	390
Prob>Chi ²	0.00	0.00	0.00	0.00	0.00
R ²	0.39	0.38	0.39	0.32	0.32

* $p < 0.1$; ** $p < 0.05$; *** $p < 0.01$

The results indicate that the relationship between credit risk and the net interest income share is nonlinear. The quadric variable of the net interest income share ratio is significant with the p values below 0.1 on the 5-year CDS maturity. The significance is even higher for the 1

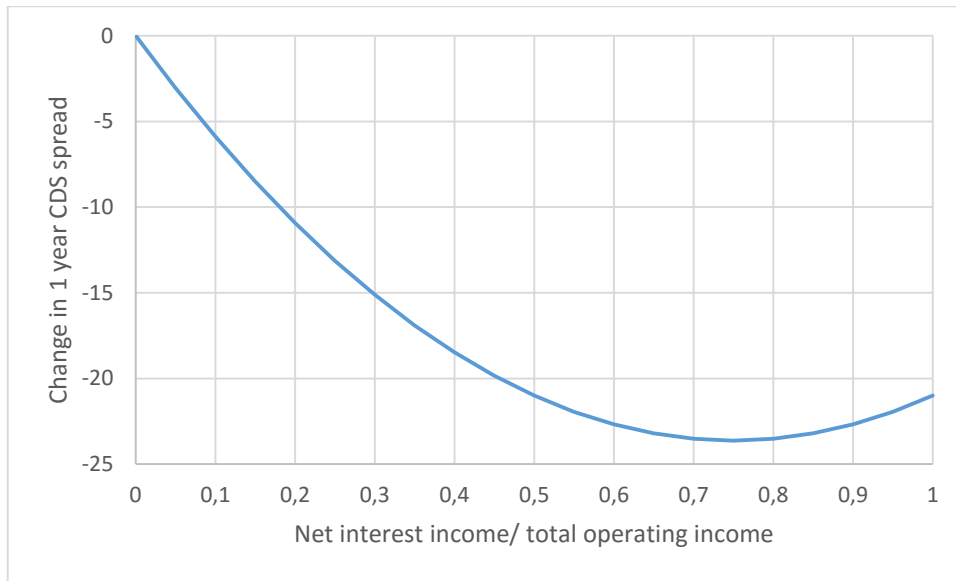
(below 0.01) and 10-year (below 0.05) maturities. The coefficient of the quadric net interest income share ratio ranges between 22 and 23 for the 5 and 10-year maturities and is 42 for the 1 year maturity. The coefficient for the linear net interest income share ratio ranges between -41 and -69. However, the linear net interest income share is not significant. Based on the coefficients' high (low) levels of non-interest income (net interest income) increases CDS spreads. The relationship is nonlinear, moving from high levels of non-interest income towards generating more net interest income reduces credit risk for the financial institutions, suggesting that investment banking is riskier than commercial banking. However, since the linear coefficient of the net interest income share is not significant, the results indicate that no significant difference based on the functional focus of the financial institution on credit risk can be determined. The significant and positive quadric term actually suggest that reductions in credit risk are due to diversification effects. Figure 12 demonstrates the change in CDS spreads relative to the level of net interest income share, using the-5 year maturity results and holding other variables constant.

Figure 12 – Non-linear Relationship between CDS Spread and Net Interest Income Share Using 5-Year maturity



Moving from generating 0% income from commercial banking to 50% offers a reduction in CDS spread of 15 basis points. When comparing the change to the mean 5 year CDS spread of 223.5 the results is somewhat economically significant. Using the 1-year maturity, the relationship is even stronger with Figure 13 describing the relationship.

Figure 13 – Non-linear Relationship between CDS Spread and Net Interest Income Share Using 1-Year maturity



In the 1-year maturity moving from generating 0% income from commercial banking to 50% offers a reduction in CDS spread of 21 basis points. The spread continues to decrease until 75% of income is generated from commercial banking, after which spread starts to increase. The increase in spread suggest that potential diversification benefits exist in the medium values of net interest income to total operating income. The relationship further suggests that the reductions in credit risk are driven by functional diversification not functional focus. Whether diversification benefits exist is further examined later in this study.

Including the quadric variable of the loan share ratio does not improve the significance of the measure, reinforcing the acceptance of the null hypothesis for the loan share.

To further examine the first research question, instead of using the financial institution level CDS spreads as the dependable credit risk variable, the difference between the financial institution level CDS spread and the banking industry CDS spread index is used. The difference is used to neutralize the effect of changes in the level of credit risk in the banking industry through time.

**Table IV – Banking Sector CDS Index Spread Differences,
Random Effects Regressions**

$Net\ interest\ income\ share = \frac{INT}{TOR}$ and $Loan\ share = \frac{LOAN}{TEA}$. The difference between the 5 year maturity financial institution level CDS spread and the 5 year banking sector CDS index is used as the dependable variable. CDS spreads are yearly averages of daily financial institution level CDS spread observations.

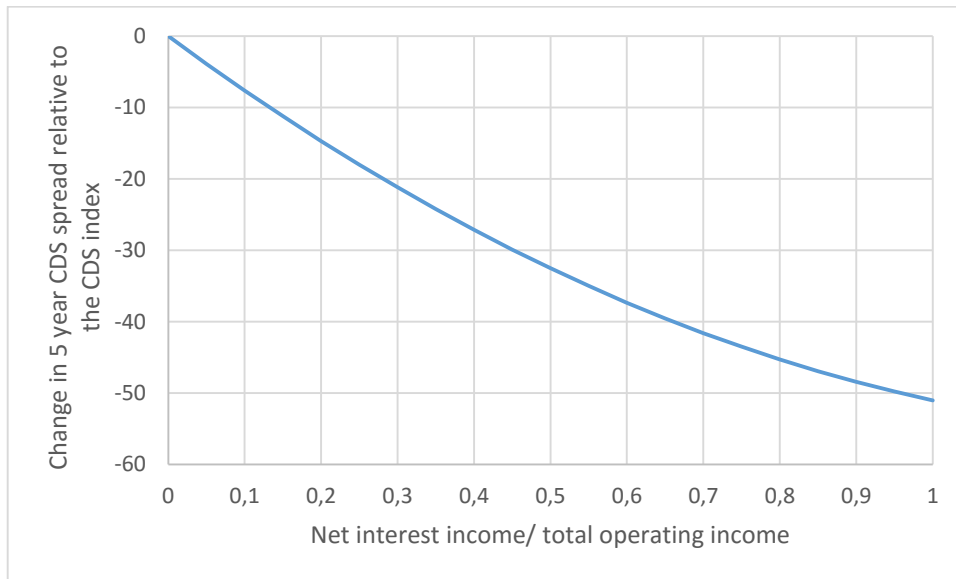
	CDS difference 5y	CDS difference 5y	CDS difference 5y
Net interest income share	-79.10 (-1.83)*		-68.14 (-1.47)
(Net interest income share)²	28.08 (2.67)***		27.08 (2.54)**
Loan share		-55.90 (-0.80)	-36.74 (-0.42)
(Loan share)²		-92.10 (-0.83)	-49.44 (-0.42)
Tier 1	600.27 <i>(1.65)*</i>	461.15 <i>(1.23)</i>	561.80 <i>(1.49)</i>
Leverage	-2941.94 <i>(-5.95)***</i>	-2810.09 <i>(-5.67)***</i>	-2851.21 <i>(-5.70)***</i>
Cost/income.	50.09 <i>(1.87)*</i>	26.01 <i>(1.45)</i>	43.73 <i>(1.53)</i>
LN(Size)	-96.74 <i>(-6.35)***</i>	-103.63 <i>(-5.37)***</i>	-99.58 <i>(-5.04)***</i>
Provisions/loans	5315.31 <i>(7.01)***</i>	5642.34 <i>(7.34)***</i>	5437.99 <i>(7.10)***</i>
Int. exp./liabilities	-502.39 <i>(-0.65)</i>	-354.39 <i>(-0.46)</i>	-464.89 <i>(-0.60)</i>
Liquid/assets	-995.64 <i>(-2.47)**</i>	-904.73 <i>(-2.21)**</i>	-962.50 <i>(-2.33)**</i>
Constant	1349.03 <i>(6.38)***</i>	1483.56 <i>(5.16)***</i>	1416.65 <i>(4.79)***</i>
N	394	394	394
Prob>Chi ²	0.00	0.00	0.00
R ²	0.34	0.35	0.35

* $p < 0.1$; ** $p < 0.05$; *** $p < 0.01$

When using the CDS spread difference, the results are stronger and more significant for the net interest income share ratio. When the loan share ratio is not used in the regression, even the linear net interest income share ratio is significant, with a p value below 0.10 and with a

coefficient of -79. Similarly the quadric net interest income share ratio is significant at below 0.01 level with a coefficient of 28. When the loan share ratio is added, the linear net interest income share is no longer significant, the quadric term remains significant at 0.05 level. Figure 14 describes the relationship using the results without the loan share ratio, holding other variables constant.

Figure 14 – Non-linear Relationship Between CDS Difference and Net Interest Income Share



The relationship is still nonlinear, however less than in previous results due to the larger linear coefficient. Moving from generating 0% income from commercial banking to 50% offers a reduction in CDS spread of 32.5 basis points compared with the banking industry CDS index. The coefficients suggest that financial institutions generating a larger share of net interest income from commercial banking are less risky than those focusing on investment banking, when comparing with the average credit risk in the banking industry. However since the coefficient of the linear net interest income share is only moderately significant in one of the regressions, the decrease in credit risk is more driven by functional diversification and no conclusion on the impact of functional focus can be drawn.

When using the CDS difference as the dependable variable for credit risk, the impact of control variables tier 1 and relative interest expense changes. As previously discussed, the relationship between the bank level CDS spreads and the tier 1 or interest expense to liabilities ratios was irrational, higher tier 1 (lower risk) and lower relative interest expense was associated

with higher financial institution level credit risk. It was previously speculated that the irrational coefficients might result from the increased CDS spreads during the financial crisis and regulatory measures in response to the crisis. When using the CDS spread difference, the impact of both ratios is much lower and less significant. The coefficient of the tier 1 ratio is lower and not significant in two of the regressions in Table IV, the coefficient remains positive. Similarly the coefficient of the interest expense to liabilities ratio remains negative, however the coefficient is much smaller and not significant in all three regressions in Table IV. Since the impact of both ratios is diminished and the coefficients are mostly not significant, the results hint that the irrational relationships with financial institution level CDS spreads were to some extent resulting from the financial crisis. Using the CDS difference controls the results for the impact of general increases in the CDS spreads in the banking sector. Thus, the effect of the financial crisis is mitigated and the results better describe the bank level heterogeneity. However, no final conclusion can be made on the impact of the two ratios and more research is called for.

It cannot be concluded that financial institutions generating a larger share of income from investment banking, are riskier in terms of credit risk than those focusing more on commercial banking, because the linear term of the net interest income share is not significant throughout the deployed regressions, with the exception of one regression. The first hypothesis of the first research question can't be accepted, the functional focus of financial institutions does not affect the credit risk of financial institutions. The results seem to be more driven by functional diversification since the quadric term of the net interest income share is significantly positive in non-linear regressions.

5.2. Are More Diversified Financial Institutions Less Risky?

This section of this study concentrates on the main focus of this study, examining if functional diversification impacts the credit risk of financial institutions. Table V presents the results using income and asset diversification measurements as the explanatory variables. Unlike the net interest income share and loan share ratios the income and asset diversification measures do not differentiate between which, commercial banking or investment banking, the financial institution focuses on. The income and asset diversification measures only measure how functionally diversified the financial institution is. However, the net interest income share and the loan share measures are used to control for the functional focus of the financial

institution, since the relationship between functional diversification and credit risk could be affected by functional focus.

Table V – Income and Asset Diversification Measures, Random Effects Regressions

$INC. DIV = 1 - \left| \frac{INT-NON.INT}{TOR} \right|$ and $AS. DIV = 1 - \left| \frac{LOAN-OEA}{TEA} \right|$. Net interest income share and loan share ratios are used as control variables. Mainly 5 year maturity CDS spread is used in the regressions as a dependable variable, for the sake of robustness also 1 and 10 year maturities are used.

	CDS 5y	CDS 5	CDS 5y	CDS 5y	CDS 1y	CDS 10y
Inc. div	-52.78 (-2.59)***	-53.58 (-2.63)***		-55.36 (-2.66)***	-94.79 (3.42)***	-55.03 (-2.98)***
As. div	-8.66 (-0.12)		-25.77 (-0.36)	-10.32 (-0.14)	3.59 (0.03)	38.77 (0.56)
Net interest income share				-17.88 (-0.38)	-12.45 (-0.20)	-42.31 (-1.02)
Loan share				67.95 (0.58)	30.54 (0.19)	61.86 (0.58)
Tier 1	1587.90 (3.93)***	1605.15 (3.96)***	1577.66 (3.88)***	1658.44 (3.95)***	1270.65 (2.25)**	1590.11 (4.23)***
Leverage	-3794.14 (-7.30)***	-3884.15 (7.37)***	-3974.13 (-7.76)***	-3863.20 (-7.30)***	-3628.36 (-4.90)***	-2528.64 (-5.29)***
Cost/income.	23.19 (1.14)	23.61 (1.17)	26.65 (1.31)	33.30 (1.04)	54.34 (1.28)	44.56 (1.57)
LN(Size)	-102.78 (-6.85)***	-104.26 (-6.90)***	-106.70 (-7.20)***	-96.35 (-4.84)***	-105.20 (-3.86)***	-80.94 (-4.41)***
Provisions/loans	7748.43 (8.96)***	7611.37 (8.87)***	7996.00 (9.22)***	7678.67 (8.84)***	9069.59 (7.72)***	5848.09 (7.47)***
Int. exp./liabilities	-1697.19 (-2.01)**	-1789.21 (-2.11)**	-1361.89 (-1.63)	-1707.72 (-2.01)**	-1459.80 (-1.28)	-1764.86 (-2.32)**
Liquid/assets	-461.38 (-1.05)	-432.15 (-0.97)	-437.22 (-0.99)	-514.46 (1.13)	-546.51 (-0.90)	-431.84 (-1.06)
Constant	1571.94 (7.66)***	1591.96 (7.56)***	1598.98 (7.85)***	1457.04 (4.82)***	1565.67 (3.79)***	1188.78 (4.27)***
N	394	394	394	394	390	390
Prob>Chi ²	0.00	0.00	0.00	0.00	0.00	0.00
R ²	0.39	0.39	0.39	0.39	0.33	0.32

* $p < 0.1$; ** $p < 0.05$; *** $p < 0.01$

The results for income diversification are promising. The coefficient for income diversification is negative and significant with p values lower than 0.01, suggesting that more income diversified financial institutions have lower CDS spreads and credit risk. The coefficient ranges between -52 and -55 for the 5 and 10-year maturities, the coefficient for the 1-year maturity is -95 and even more significant. When controlling for the functional focus of the financial institution, commercial banking or investment banking, the coefficient of income diversification is not largely affected. Thus the results are driven by diversification and no evidence can be obtained that functional focus impacts credit risk. Using the -55 coefficient of 5-year maturity, an increase of income diversification from the first quartile of 0.532 to the third quartile of 0.891 decreases CDS spreads by 19.7 basis points. With the -95 coefficient of 1-year maturity, the same effect corresponds to a decrease of 35.5 basis points. Compared with the 223.5 basis point mean of the 5-year CDS spreads and the 186.7 mean of the 1-year maturity, both results are economically meaningful.

The results for the asset diversification measure are not significant across the different regressions and the sign of the asset diversification changes for different maturities. Thus the null hypothesis can't be rejected and it is concluded that the asset diversification does not impact the credit risk of financial institutions. The asset diversification is not a fully representative measure for functional diversification, since especially investment banking is not fully represented in the balance sheet. Services like investment banking do not require large assets since income is mostly generated through fees from advisory and underwriting services. Moreover, investment banking generates assets outside the balance sheet which are not captured by the assets diversification measure. Income based diversification measures capture diversification to a larger extent and the non-significant asset diversification results do not take away or contradict the results gained by the income diversification.

Compared with the results using the net interest income and loan share ratios in tables II and III, the control variables have the same signs and similar coefficients with mostly the same significance.

In order to test if the relationship between the credit risk and functional diversification is nonlinear, the linear and quadric income and asset diversification measures are used as explanatory variables. Table VI presents results adding the quadric diversification variables and the better fitting quadric control variable for the net interest income share.

Table VI – Income and Asset Diversification, Random Effects Regressions Using Quadric Terms

Mainly 5 year maturity CDS spread is used in the regressions as a dependable variable, for the sale of robustness also 1 and 10 year maturities are used.

	CDS 5y	CDS 5y	CDS 5y	CDS 5y	CDS 1y	CDS 10y
Inc. div	-85.13 (-2.40)**	-85.21 (-2.50)**	2490.43 (0.30)	-105.55 (-2.17)**	-160.04 (2.50)**	-98.76 (-2.30)**
(Inc. div)²	-7.90 (-1.17)	-7.57 (-1.17)	-1297.92 (-0.31)			
As. div	-64.30 (-0.24)		-83.32 (-0.30)	-86.50 (-0.32)	20.98 (0.21)	56.23 (0.82)
(As. div)²	61.44 (0.27)		73.13 (0.32)	73.71 (0.33)		
Net interest income share			-5139.17 (-0.31)	53.23 (0.87)	58.68 (0.71)	11.02 (0.20)
(Net interest income share) ²			5156.55 (0.31)	-35.33 (-1.25)	-42.67 (-1.14)	-29.20 (-1.17)
Tier 1	1514.42 (3.73)***	1558.00 (3.85)***	1494.42 (3.66)***	1489.90 (3.66)***	1177.61 (2.16)**	1495.16 (4.10)***
Leverage	-3616.01 (-6.91)***	-3725.59 (-7.07)***	-3579.93 (-6.86)***	-3563.33 (-6.87)***	-3428.29 (-4.67)***	-2475.77 (4.99)***
Cost/income.	24.19 (1.18)	24.10 (1.19)	15.25 (0.49)	15.07 (0.48)	40.89 (0.98)	32.45 (1.17)
LN(Size)	-101.83 (-6.89)***	-102.39 (-6.90)***	-100.02 (-6.67)***	-100.07 (-6.76)***	-105.95 (-5.07)***	-86.18 (-6.05)***
Provisions/loans	7796.48 (8.79)***	7601.19 (8.82)***	7863.42 (8.83)***	7884.71 (8.88)***	9029.89 (7.67)***	5805.64 (7.40)***
Int. exp./liabilities	-1737.72 (-2.06)**	-1833.42 (-2.16)**	-1757.63 (-2.07)**	-1727.37 (-2.04)**	-1624.75 (-1.42)	-1880.75 (-2.45)**
Liquid/assets	-526.58 (-1.20)	-491.38 (-1.11)	-528.07 (-1.20)	-534.77 (-1.22)	-571.99 (-0.97)	-415.61 (-1.05)
Constant	1593.07 (7.15)***	1591.84 (7.69)***	1575.65 (7.06)***	1576.17 (7.14)***	1613.30 (5.68)***	1301.42 (6.74)***
N	394	394	394	394	390	390
Prob>Chi ²	0.00	0.00	0.00	0.00	0.00	0.00
R ²	0.40	0.40	0.41	0.41	0.34	0.33

* $p < 0.1$; ** $p < 0.05$; *** $p < 0.01$

When including the quadric term of the income diversification measure in the regressions, the coefficient of the linear income diversification measure increases to -85. However, the significance of the linear income diversification measure decreases, with the coefficient now significant at the below 0.05 p value level. The coefficient of the quadric income diversification measure is also negative, however not significant. Since the significance of the linear term decreases and the quadric term is insignificant, it is concluded that the relationship between the CDS spreads and the functional diversification is linear. The conclusion is further enhanced when both the linear and quadric terms of the net interest income share are added to the regressions, with the linear and quadric terms of the income diversification. Coefficients of the income diversification increase dramatically and change to insignificant, since both the quadric and linear net interest income shares capture parts of the diversification impact. Also since both income diversification and net interest income share variables are derived from the same income statement data, the results are subject to multicollinearity due to high correlation between the variables. Multicollinearity can decrease the significance and cause erratic coefficient estimates. Thus for the income diversification measure only the linear term is later used. Furthermore, in regressions using the linear and quadric terms of the net interest income share as control variables without the quadric income diversification term, the coefficient of the linear income diversification measure increases.

The coefficient of income diversification for the 1 and 10-year maturities is -106 and -99 respectively, for the 1 year maturity the coefficient is -160. All of the coefficients are significant at the below 0.05 levels. Functional diversification decreases the credit risk of financial institutions, even when the nonlinear specification of the net interest income share is used as a control variable to control for the functional focus of the financial institution. Using the -106 coefficient of the 5-year maturity, an increase of income diversification from the first quartile of 0.532 to the third quartile of 0.891 decreases CDS spreads by 38.1 basis points. Using the 1-year maturity coefficient of -160, the spread reduction of the same change in income diversification amounts to -57.4 basis points.

The results further suggest that the significant quadric net interest income share results in the first research question were driven by functional diversification, not functional focus. Since in Table VI both the linear and quadric terms of the net interest income share are not significant and income diversification captures the diversification impact.

To further test the impact of functional diversification, Table VII shows results using the difference between the financial institution level CDS spreads and the banking industry CDS index spread as the dependable variable for credit risk.

Table VII – Income and Asset Diversification, Random Effects Regressions using banking Sector CDS Index Spread Differences

$INC.DIV = 1 - \left| \frac{INT-NON.INT}{TOR} \right|$ and $AS.DIV = 1 - \left| \frac{LOAN-OEA}{TEA} \right|$. The difference between the 5 year maturity financial institution level CDS spread and the 5 year banking sector CDS index is used as the dependable variable. CDS spreads are yearly averages of daily CDS spread observations.

	CDS difference 5y	CDS difference 5y	CDS difference 5y	CDS difference 5y	CDS difference 5y
Inc. div	-59.64 (-3.29)***		-58.77 (-3.23)***	-93.38 (2.18)**	-92.84 (-2.20)**
As. div		-50.92 (-0.74)	-30.03 (-0.43)	-4.08 (-0.06)	
Net interest income share				-1.85 (-0.03)	-3.49 (-0.06)
(Net interest income share) ²				-20.92 (-0.85)	-20.42 (-0.84)
Tier 1	561.27 (1.55)	530.40 (1.45)	560.26 (1.55)	553.46 (1.52)	55.98 (1.53)
Leverage	-2754.29 (-5.56)***	-2859.69 (-5.90)***	-2679.92 (-5.48)***	-2636.46 (-5.25)***	-2662.33 (-5.30)***
Cost/income.	22.69 (1.30)	27.01 (1.51)	22.13 (1.25)	34.83 (1.25)	35.48 (1.30)
LN(Size)	-91.19 (-6.06)***	-92.95 (-6.32)***	-88.79 (-5.95)***	-90.57 (-5.89)***	-91.16 (-6.01)***
Provisions/loans	5266.10 (6.98)***	5692.76 (-7.41)***	5414.60 (7.11)***	5249.41 (6.85)***	5211.04 (6.88)***
Int. exp./liabilities	-684.57 (-0.89)	-331.82 (-0.43)	-624.77 (-0.81)	-682.92 (-0.88)	-696.83 (-0.90)
Liquid/assets	-1007.02 (-2.51)**	-965.15 (-2.40)**	-1010.55 (-2.53)**	-1045.39 (-2.60)***	-1045.08 (-2.60)***
Constant	1296.26 (6.26)***	1307.05 (6.53)***	1276.88 (6.32)***	1309.03 (6.33)***	1315.82 (6.32)***
N	394	394	394	394	394
Prob>Chi ²	0.00	0.00	0.00	0.00	0.00
R ²	0.36	0.36	0.37	0.37	0.37

* $p < 0.1$; ** $p < 0.05$; *** $p < 0.01$

When using the CDS difference as the dependable variable, the explanatory variable income diversification remains negative and significant. The asset diversification remains insignificant and the null hypothesis cannot be rejected for the asset diversification. The coefficient of the income diversification measure ranges between -60 and -59 with p values below 0.01, when the net interest income share is not used as a control variable. When the net interest income share is added as a control variable, the coefficient of the income diversification measure increases to -93, however the significance of the coefficient decreases with p values below 0.05 level. Using the -60 coefficient, an increase of income diversification from the first quartile of 0.532 to the third quartile of 0.891, decreases CDS spreads by 21.5 basis points. A similar change in diversification decreases CDS spreads by 33.4 basis points when the -93 coefficient is used. Overall, better functionally diversified financial institutions have lower credit risk when compared with the average credit risk in the banking industry.

The results in Table VII confirm that the irrational relationship between the credit risk of financial institutions and the control variables tier 1 and interest expense share, could result from the financial crisis. Again for both control variables the signs of the coefficients remain irrational, high tier 1 (low risk) increases credit risk and high interest expense reduces credit risk. However, the coefficients are insignificant when the CDS difference is used, suggesting that both results are driven by the increased overall credit risk level in the banking industry during the observation period.

In conclusion, income diversification is negatively associated with CDS spreads. The results show that financial institutions with higher income diversification have lower CDS spreads at the financial institution level and compared with the banking industry CDS index. The relationship between income diversification and CDS spreads is linear. The negative relationship between diversification and credit risk remains even when the functional focus of the financial institution is controlled for. When using the 5 or 10-year CDS maturities the coefficients of the income diversification measures are somewhat similar. With the 1-year maturity the coefficient is almost twice as high, suggesting that diversification benefits decrease the probability of default of financial institutions more in the short term.

6. Robustness

In order to confirm the gained results several robustness tests are conducted by controlling for time effects, unobserved heterogeneity, endogeneity, heteroscedasticity and using alternative functional diversification measures. Furthermore, results are confirmed by using different data samples, excluding years affected by the financial crisis from the sample and without US and Southern European financial institutions. The results remain robust, financial institutions with higher functional diversification have lower credit risk and no difference in credit risk is discovered based on functional focus.

6.1. Fixed Effects Panel Data Regressions

In order to control for unobserved financial institution level heterogeneity and potential time effects, the financial institution and year effects has been fixed using dummy variables. The fixed effect regressions should control for the possibility that the results are driven by the financial crisis or other time effect or potential endogeneity and the unobserved characteristics of financial institutions.

6.1.1. Is Investment Banking More Risky Than Commercial Banking

Table VIII – Net Interest Income Share and Loan Share, Financial Institution and Time Fixed Effects Regressions

$Net\ interest\ income\ share = \frac{INT}{TOR}$ and $Loan\ share = \frac{LOAN}{TEA}$. In the regressions the unobserved financial institution level heterogeneity is controlled for by using dummy variables for the financial institutions. Also time effect is controlled for by using year dummy variables. The results for the financial institution and year dummy variables are omitted from the table. Mainly 5 year maturity CDS spread is used in the regressions as a dependable variable, for the sake of robustness also 1 and 10 year maturities are used. CDS spreads are yearly averages of daily CDS spread observations.

	CDS 5y	CDS 5y	CDS 5y	CDS 1y	CDS 10y	CDS diff.
Net interest income share	-1.19 (-0.03)	-87.45 (-1.95)*	-87.64 (-1.96)*	-95.62 (-1.41)	-83.85 (-2.17)**	-98.55 (-2.20)**
(Net interest income share)²		41.67 (3.98)***	41.47 (3.98)***	61.92 (3.92)***	35.33 (3.91)***	43.46 (4.15)***
Loan share	-33.14 (-0.15)	-52.88 (-0.08)		-243.13 (-0.73)	-147.37 (-0.77)	-154.70 (-0.70)
(Loan share) ²		-26.31 (-0.05)				
Tier 1	12.53 (0.03)	256.30 (0.55)	273.68 (0.59)	-112.98 (-0.16)	-89.88 (-0.22)	-123.16 (-0.26)
Leverage	-5482.17 (-7.59)***	-5613.28 (-7.89)***	-5635.70 (-8.02)***	-5623.27 (-4.87)***	-3426.56 (-5.19)***	-5081.18 (-7.18)***
Cost/income.	33.01 (1.18)	49.00 (1.77)*	49.67 (1.81)*	67.68 (1.62)	48.44 (2.03)**	55.63 (2.01)**
LN(Size)	-234.29 (-4.37)***	-245.69 (-4.66)***	-239.77 (-4.81)***	-351.27 (-4.42)***	-233.29 (-5.14)***	-214.94 (-4.09)***
Provisions/loans	5103.47 (5.99)***	4859.61 (5.73)***	4928.28 (6.03)***	5538.59 (4.30)***	4034.37 (5.47)***	3551.16 (4.25)***
Int. exp./liabilities	2234.93 (1.96)*	2348.20 (2.10)**	2371.41 (2.13)**	4067.54 (2.37)**	2561.15 (2.61)***	1866.78 (1.67)*
Liquid/assets	-1579.83 (-3.23)***	-1598.35 (-3.33)***	-1654.84 (-3.64)***	-1604.76 (-2.21)**	-1336.15 (-3.22)***	-1821.76 (-3.80)***
N	394	394	394	390	390	394
Prob>F	0.00	0.00	0.00	0.00	0.00	0.00
R ²	0.76	0.77	0.77	0.66	0.76	0.71

* $p < 0.1$; ** $p < 0.05$; *** $p < 0.01$

When using the fixed effects regressions the linear net interest income share increases in significance and remains negative. The quadric term remains highly significant and positive.

Only when the 1-year CDS maturity is used, the linear net interest income share is insignificant. Based on the results, financial institutions with a higher share of non-interest income have higher credit risk when individual characteristics and time effects are controlled for. The relationship is non-linear, reductions in credit risk are higher for financial institutions moving from generating high amounts of non-interest income towards generating more net interest income. The fixed effects regressions improve results for the first research question and suggest that financial institutions focusing more on investment banking have higher credit risk. However, the results are not particularly strong in significance and later in the robustness section it is discovered that the functional diversification effect impacts the result to a higher degree. Thus the first hypothesis of the first research question can't be accepted and no difference in credit risk is discovered between financial institutions focusing more on commercial banking or investment banking.

The use of fixed effect regressions had an impact on the coefficients of the control variables tier 1, cost to income, interest expense to liabilities and liquid assets to total assets. In previous regressions, the coefficient of tier 1 was positive and significant, suggesting that higher tier 1 ratio (lower risk) would result in higher credit risk. Also the coefficient of the interest expense to total liabilities had an illogical relationship, higher relative interest expense decreased CDS spreads. As previously hypothesized, the time effect caused by the financial crisis could have impacted the results and created the illogical coefficients. When time effects are controlled for, the coefficient for tier 1 is no longer significant and even negative when using the 1 and 10-year maturity CDS spreads or the CDS difference. The coefficients for the interest expense share are positive and significant, suggesting that higher relative interest expense increases credit risk as expected when time effect is controlled for.

In previous regressions, the cost to income and liquid assets to total assets ratios were insignificant in most of the regressions, some significance was gained when the CDS difference was used as the dependable variable. However, when the fixed effects regressions are used, both variables gained significance. Cost to income ratio is positive, suggesting that less profitable financial institutions have higher credit risk, however the coefficient is only significant in some of the regressions. The liquid assets to total assets ratio is highly significant in all of the regressions at below 0.01 level. The coefficient is negative, suggesting that higher relative amount of liquid assets reduces credit risk for financial institutions.

6.1.2. Are More Diversified Financial Institutions Less Risky?

When using the income and asset diversification measures in the financial institution and year fixed effects regressions the results remain robust. The coefficients of the income diversification remain negative and highly significant. Asset diversification on the other hand remains insignificant. Even when controlling for the unobserved individual characteristics of financial institutions and time effects, financial institutions with higher functional diversification have lower credit risk. Table IX presents the results.

Table IX – Income and Asset Diversification, Financial Institution and Time Fixed Effects Regressions

$INC.DIV = 1 - \left| \frac{INT-NON.INT}{TOR} \right|$ and $AS.DIV = 1 - \left| \frac{LOAN-OEA}{TEA} \right|$. In the regressions the unobserved financial institution level heterogeneity is controlled for, by using dummy variables for the financial institutions. Also time effect is controlled for by using year dummy variables. The results for the financial institution and year dummy variables are omitted from the table. Mainly 5 year maturity CDS spread is used in the regressions as a dependable variable, for the sake of robustness also 1 and 10 year maturities are used. CDS spreads are yearly averages of daily CDS spread observations.

	CDS 5y	CDS 5y	CDS 5y	CDS 1y	CDS 10y	CDS diff.
Inc. div	-73.56 (-3.85)***	-75.15 (-3.89)***	-30.81 (-0.68)	-118.28 (-4.06)***	-65.90 (-3.95)***	-79.83 (-4.13)***
As. div	52.97 (0.46)	61.02 (0.52)	48.75 (0.42)	148.50 (0.84)	65.44 (0.65)	69.39 (0.60)
Net interest income share		-23.51 (-0.59)	-66.56 (1.19)	-5.74 (-0.10)	-31.54 (-0.92)	-33.29 (-0.84)
(Net interest income share) ²			26.57 (1.09)			
Tier 1	302.60 (0.65)	336.17 (0.72)	313.35 (0.67)	72.23 (0.10)	10.27 (0.03)	-15.20 (-0.03)
Leverage	-5419.67 (-7.69)***	-5428.90 (-7.69)***	-5568.01 (-7.77)***	-5398.91 (-4.72)***	-3297.75 (-5.04)***	-4915.13 (-6.96)***
Cost/income.	25.02 (1.48)	37.75 (1.38)	46.09 (1.62)	53.81 (1.30)	40.16 (1.70)*	44.88 (1.64)
LN(Size)	-236.33 (-4.72)***	-236.15 (-4.72)***	-239.35 (-4.77)***	-329.35 (-4.36)***	-219.60 (-5.07)***	-199.92 (-3.99)***
Provisions/loans	4937.21 (5.97)***	4872.09 (5.84)***	4862.28 (5.82)***	5565.17 (4.33)***	4084.39 (5.54)***	3598.30 (4.31)***
Int. exp./liabilities	2352.12 (2.11)**	2342.88 (2.10)**	2373.93 (2.13)**	4104.62 (2.40)**	2577.56 (2.63)***	1889.90 (1.69)*
Liquid/assets	-1616.33 (-3.52)***	1609.05 (-3.50)***	-1625.04 (-3.53)***	-16751.68 (-2.41)**	-1392.71 (-3.49)***	-1876.26 (-4.08)***
N	394	394	394	390	390	394
Prob>F	0.00	0.00	0.00	0.00	0.00	0.00
R ²	0.77	0.77	0.77	0.66	0.76	0.71

* $p < 0.1$; ** $p < 0.05$; *** $p < 0.01$

Controlling for financial institution fixed effects alleviates endogeneity concerns to some extent. Financial institutions choose to functionally diversify based on unknown financial

institution level decision variables or characteristics. The same characteristics could drive the CDS spreads determined by investors in the CDS market, instead of income diversification itself as suggested by Laeven and Levine (2007). Without controlling for the unobserved characteristics, biased results for the impact of income diversification could be obtained. By controlling for the unobserved characteristics with financial institution dummies, the endogeneity concerns should be alleviated. Also by controlling for financial institutions level unobserved characteristics, the concern that the impact of income diversification is caused by the difference in the prospects or abilities of financial institutions to expand into new activities is alleviated.

The results further suggest that the significant linear and quadric net interest income share results, in the fixed effects regressions for the first research question, were driven by functional diversification, not functional focus. Since in table IX both the linear and quadric terms of the net interest income share are not significant and income diversification captures the diversification impact.

6.2. Heteroscedasticity

Heteroscedasticity (non-constant variance of the disturbance term) can cause bias in the estimates for the standard errors of the regression coefficients, affecting the significance of the coefficients. The actual coefficients are not biased by heteroscedasticity but since the significance tests are, heteroscedasticity can lead to wrong interpretations of the coefficients and the null hypothesis can be wrongly rejected. Using the modified Wald test for groupwise heteroscedasticity for fixed effects regressions, heteroscedasticity between the financial institution groups was found at below 0.00 F-test level. Heteroscedasticity was also detected for the random effects regressions using the likelihood ratio test at below 0.00 F-test level. Since heteroscedasticity is observed, the GLS heteroscedastic robust standard errors method is used. The method adjusts the standard errors for all financial institution groups, to take heteroscedasticity into account. GLS method adjusts the standard errors for heteroscedasticity and should provide non-biased significance test results. However, only the use of linear variables is allowed. The GLS method is only used for the income and asset diversification regressions, since no significant results were obtained for the linear net interest income share and loan share regressions.

Table X – GLS Heteroscedastic Robust Standard Error Regressions

	Random Effects			Financial Institution and Year Fixed Effects		
	CDS 5y	CDS 5y	CDS diff.	CDS 5y	CDS 5y	CDS diff.
Inc. div	-52.79 <i>(-2.60)***</i>	-53.22 <i>(-2.63)***</i>	-60.34 <i>(-3.70)***</i>	-73.56 <i>(-2.78)***</i>	-75.15 <i>(-2.75)***</i>	-79.83 <i>(-3.05)***</i>
As. div	-8.66 <i>(-0.10)</i>	-3.94 <i>(-0.05)</i>	-14.50 <i>(-0.17)</i>	52.97 <i>(0.49)</i>	61.02 <i>(0.55)</i>	69.39 <i>(0.62)</i>
Net interest income share		-8.63 <i>(-0.16)</i>	-33.99 <i>(-0.81)</i>		-23.51 <i>(-0.47)</i>	-33.29 <i>(-0.67)</i>
Tier 1	1587.90 <i>(1.98)**</i>	1594.19 <i>(2.00)**</i>	576.99 <i>(0.82)</i>	302.60 <i>(0.52)</i>	336.17 <i>(0.59)</i>	-15.20 <i>(-0.03)</i>
Leverage	-3794.14 <i>(-2.39)**</i>	-3810.95 <i>(-2.38)**</i>	-2699.55 <i>(-1.63)</i>	-5419.67 <i>(-3.52)***</i>	-5428.90 <i>(-3.52)***</i>	-4915.13 <i>(-3.17)***</i>
Cost/income.	23.19 <i>(0.72)</i>	27.69 <i>(0.62)</i>	39.95 <i>(1.03)</i>	25.02 <i>(1.05)</i>	37.75 <i>(1.25)</i>	44.88 <i>(1.47)</i>
LN(Size)	-102.78 <i>(-4.45)***</i>	-103.59 <i>(-4.64)***</i>	-91.35 <i>(-3.57)***</i>	-236.33 <i>(-4.19)***</i>	-236.15 <i>(-4.20)***</i>	-199.92 <i>(-3.65)***</i>
Provisions/loans	7748.43 <i>(4.34)***</i>	7712.02 <i>(4.30)***</i>	5330.43 <i>(3.51)***</i>	4937.21 <i>(3.59)***</i>	4872.09 <i>(3.61)***</i>	3598.90 <i>(2.61)***</i>
Int. exp./liabilities	-1697.19 <i>(-1.92)*</i>	-1704.28 <i>(-1.91)*</i>	-601.74 <i>(-0.67)</i>	2352.12 <i>(1.94)*</i>	2342.88 <i>(1.93)*</i>	1889.89 <i>(1.65)*</i>
Liquid/assets	-461.38 <i>(-0.94)</i>	-461.63 <i>(-0.95)</i>	-1031.03 <i>(-2.79)***</i>	-1616.33 <i>(-3.78)***</i>	-1609.05 <i>(-3.73)***</i>	-1876.25 <i>(-4.14)***</i>
Constant	1571.94 <i>(4.04)***</i>	1582.41 <i>(4.13)***</i>	1308.33 <i>(3.18)***</i>			
N	394	394	394	394	394	394
Prob>Chi ² / Prob>F	0.00	0.00	0.00	0.00	0.00	0.00
R ²	0.39	0.39	0.36	0.77	0.77	0.71

* $p < 0.1$; ** $p < 0.05$; *** $p < 0.01$

Based on the GLS results, the coefficients for the income diversification are negative and highly significant with p values lower than 0.01 in all of the regressions using random and fixed effects. The results indicate that previous results were not driven by heteroscedasticity, since with robust standard errors the coefficients are highly significant. The results of the GLS regressions further suggest that financial institutions with higher functional diversification have lower CDS spreads.

6.3. Results Using the Herfindahl-Hirschman Diversification Measures

To further test the impact of functional diversification on the credit risk of financial institutions, alternative income and asset diversification measures using the modified Herfindahl-Hirschman index are employed. Table XI presents the results.

**Table XI – HH Income and Asset Diversification Measures, Random Effects Panel
Data Regressions**

$HH\ INC.\ DIV = 1 - \left(\left(\frac{INT}{TOR} \right)^2 + \left(\frac{NON.INT}{TOR} \right)^2 \right)$ and $HH\ AS.\ DIV = 1 - \left(\left(\frac{LOAN}{TEA} \right)^2 + \left(\frac{OEA}{TEA} \right)^2 \right)$. Net interest income share and loan share ratios are used as control variables. Mainly 5 year maturity CDS spread is used in the regressions as a dependable variable, for the sake of robustness also 1 and 10 year maturities are used. CDS spreads are yearly averages of daily CDS spread observations.

	CDS 5y	CDS 5y	CDS 5y	CDS 5y	CDS 1y	CDS 10y
HH inc. div	-10.32 (-1.76)*	-10.62 (-1.82)*		-11.14 (-1.84)*	-20.82 (-2.59)***	-11.51 (-2.15)**
HH as. div.	-104.81 (-0.68)		-97.90 (-0.64)	-122.82 (-0.72)	-150.48 (-0.64)	11.87 (0.08)
Net interest income share				-19.05 (-0.4)	-21.44 (-0.34)	-46.38 (-1.09)
Loan share				64.81 (0.55)	27.94 (0.17)	46.75 (0.43)
Tier 1	1590.76 (3.93)***	1619.23 (3.98)***	1561.38 (3.85)***	1659.42 (3.94)***	1293.94 (2.28)**	1590.65 (4.21)***
Leverage	-3873.50 (-7.62)***	-4011.01 (-7.68)***	-3927.47 (-7.75)***	-3941.20 (-7.57)***	-3859.52 (-5.27)***	-2727.96 (-5.56)***
Cost/income.	23.21 (1.13)	24.87 (1.22)	26.26 (1.28)	33.65 (1.05)	58.64 (1.37)	46.34 (1.63)
LN(Size)	-103.24 (-7.16)***	-106.82 (-7.13)***	-105.58 (-7.38)***	-97.00 (-4.94)***	-107.15 (-3.95)***	-83.19 (-4.56)***
Provisions/loans	8091.16 (9.30)***	7769.14 (9.03)***	8142.81 (9.34)***	8047.84 (9.21)***	9508.25 (8.03)***	6085.71 (7.71)***
Int. exp./liabilities	-1413.73 (-1.7)*	-1594.25 (-1.88)*	-1291.56 (-1.56)	-1413.00 (-1.69)*	-1109.37 (-0.98)	-1520.91 (-2.01)**
Liquid/assets	-447.08 (-1.02)	-394.92 (-0.89)	-456.35 (-1.04)	-500.41 (-1.11)	-494.37 (-0.82)	-414.19 (-1.03)
Constant	1578.72 (7.74)***	1582.01 (7.57)***	1605.26 (7.91)***	1472.81 (5.01)***	1601.03 (3.95)***	1212.44 (4.45)***
N	394	394	394	394	390	390
Prob>Chi ²	0.00	0.00	0.00	0.00	0.00	0.00
R ²	0.39	0.38	0.39	0.39	0.32	

* $p < 0.1$; ** $p < 0.05$; *** $p < 0.01$

The HH income diversification results are in line with the previously obtained results using the normal income diversification method. The coefficient of the HH income diversification is negative, financial institutions with higher functional diversification have lower CDS spreads and credit risk. However, the coefficients and the economic impacts are lower for the HH income diversification. The coefficient ranges between -10 and -12 for the 5 and 10-year maturities and -21 for the 1-year maturity. Using the -11 coefficient of the 5-year maturity, an increase of HH income diversification from the first quartile of 0.390 to the third quartile of 0.494 decreases CDS spreads by only 1.0 basis points. Furthermore, the coefficients of HH income diversification are less significant than the regular income diversification measure, only less than 0.1 p values can be obtained for the 5-year maturity coefficients. For the 1 and 10-year maturities the significance is higher at below 0.01 and 0.05 levels respectively. However, interpreting the Herfindahl-Hirschman index is less straightforward than the regular income diversification measure, as the HH index uses the sum of squares of the net interest income to total operating income and non-interest income to total operating income ratios. Thus for economic impact the regular income diversification measure is primarily used and the HH income diversification acts as a way to confirm that functional diversification reduces credit risk. As with previous results the HH asset diversification coefficient is insignificant and the null hypothesis can't be rejected.

In Table XIII it is tested if the relationship between CDS spreads and the HH income diversification measures are non-linear. The quadric terms of the HH income measures are added to the regressions in Table XII.

**Table XII – HH Income and Asset Diversification Measures, Random Effects Panel
Data Regressions Using Quadric Terms**

$HH\ INC.DIV = 1 - \left(\left(\frac{INT}{TOR} \right)^2 + \left(\frac{NON.INT}{TOR} \right)^2 \right)$ and $HH\ AS.DIV = 1 - \left(\left(\frac{LOAN}{TEA} \right)^2 + \left(\frac{OEA}{TEA} \right)^2 \right)$. Net interest income share and loan share ratios are used as control variables. Mainly 5 year maturity CDS spread is used in the regressions as a dependable variable, for the sake of robustness also 1 and 10 year maturities are used. CDS spreads are yearly averages of daily CDS spread observations.

	CDS 5y	CDS 5y	CDS 5y	CDS 5y	CDS 1y	CDS 10y
HH inc. div	-64.25 (-2.57)***	-65.44 (-2.64)***	2296.38 (0.28)	-89.32 (-2.92)***	-148.08 (-3.67)***	-86.79 (-3.22)***
(HH inc. div)²	-1.98 (-2.21)**	-2.02 (-2.28)**	-2.98 (-2.61)***	-2.99 (-2.62)***	-4.87 (-3.23)***	-2.88 (2.87)***
HH as. div.	-59.21 (-0.38)		-116.04 (-0.72)	-120.49 (-0.75)	-219.98 (-1.06)	-6.34 (-0.04)
Net interest income share			-4689.29 (-0.28)	82.32 (1.43)	140.34 (1.85)*	52.30 (1.02)
(Net interest income share) ²			4771.21 (0.29)			
Tier 1	1562.57 (3.87)***	1584.52 (3.91)***	1506.10 (3.72)***	1502.68 (3.71)***	1080.94 (2.01)**	1447.71 (4.00)***
Leverage	-3760.65 (-7.32)***	-3856.68 (-7.38)***	-3703.13 (-7.19)***	-3688.08 (-7.20)***	-3385.37 (-4.91)***	-2457.87 (-5.18)***
Cost/income.	39.54 (1.82)*	40.82 (1.91)*	6.15 (0.19)	6.09 (0.19)	15.94 (0.38)	18.97 (0.67)
LN(Size)	-103.97 (-7.19)***	-106.21 (-7.14)***	-99.16 (-6.65)***	-99.24 (-6.74)***	-99.42 (-5.18)***	-82.73 (-6.15)***
Provisions/loans	7750.81 (8.86)***	7540.05 (8.75)***	7732.07 (8.82)***	7747.78 (8.86)***	9365.99 (7.97)***	5903.49 (7.51)***
Int. exp./liabilities	-1750.81 (-2.03)**	-1826.47 (-2.15)**	-1946.78 (2.27)**	-1920.63 (-2.25)**	-1798.90 (-1.61)	-1950.12 (-2.57)*
Liquid/assets	-440.33 (-1.01)	-405.00 (-0.92)	-384.28 (-0.88)	-388.38 (-0.89)	-411.57 (-0.72)	-347.89 (-0.89)
Constant	1585.72 (7.72)***	1597.19 (7.65)***	1540.85 (7.41)***	1542.14 (7.50)***	1543.41 (5.74)***	1241.66 (6.62)***
N	394	394	394	394	390	390
Prob>Chi ²	0.00	0.00	0.00	0.00	0.00	0.00
R ²	0.40	0.39	0.41	0.41	0.36	0.34

* $p < 0.1$; ** $p < 0.05$; *** $p < 0.01$

Including the quadric term of the HH income diversification in the regressions increases the significance of the linear HH income diversification. The linear term is now highly significant with below 0.01 p values. Only when the quadric term of the net interest income share is added as a control variable, the linear HH income diversification is insignificant potentially due to multicollinearity. The quadric term of the HH income diversification is significant below the 0.01 level only when the linear net interest income share is used as a control variable. Also the economic impact increases as the coefficients are higher. When using the coefficients of -89 for the linear HH income diversification and -3 for the quadric HH income diversification, an increase of HH income diversification from the first quartile of 0.390 to the third quartile of 0.494 decreases CDS spreads by 9.5 basis points.

Table XIII describes the results using the difference between the financial institution level CDS spreads and the banking sector CDS index as the dependable variable.

Table XIII – HH Income and Asset Diversification Measures, Random Effects Panel Data Regressions Using Banking Sector CDS Index Spread Differences

The difference between the 5 year maturity financial institution level CDS spread and the 5 year banking sector CDS index is used as the dependable variable. CDS spreads are yearly averages of daily CDS spread observations.

	CDS difference 5y	CDS difference 5y	CDS difference 5y	CDS difference 5y
HH inc. div	-58.48 (-2.69)***	-12.18 (-2.36)**	3043.29 (0.41)	-65.71 (-2.46)**
(HH inc. div)²	-1.69 (-2.18)**		-1.98 (-1.99)**	-1.98 (-1.99)**
HH as. div.	-128.67 (-0.85)	-168.53 (-1.12)	-139.55 (-0.89)	-145.08 (-0.93)
Net interest income share			-6193.95 (-0.42)	24.13 (0.47)
(Net interest income share) ²			6217.83 (0.42)	
Tier 1	541.66 (1.50)	562.20 (1.55)	527.04 (1.45)	524.75 (1.44)
Leverage	-2689.06 (-5.55)***	-2786.12 (-5.80)***	-2689 (-5.50)	-2671.60 (-5.49)***
Cost/income.	36.50 (1.94)*	22.60 (1.27)	26.40 (0.92)	26.45 (0.92)
LN(Size)	-90.26 (-6.24)***	-89.66 (-6.24)***	-88.83 (-5.95)***	-88.93 (-6.03)***
Provisions/loans	5472.76 (7.10)***	5754.72 (7.50)***	5453.31 (7.06)***	5472.06 (7.09)***
Int. exp./liabilities	-643.96 (-0.84)	-420.16 (-0.55)	-734.53 (-0.94)	-707.80 (-0.91)
Liquid/assets	-970.65 (-2.44)**	-964.22 (-2.42)**	-959.51 (-2.40)**	-956.38 (-2.39)**
Constant	1304.96 (6.45)***	1300.28 (6.46)***	1291.44 (6.26)***	1292.64 (6.33)***
N	394	394	394	394
Prob>Chi ²	0.00	0.00	0.00	0.00
R ²	0.37	0.36	0.37	0.37

* $p < 0.1$; ** $p < 0.05$; *** $p < 0.01$

The results are consistent with previous results, significance and the coefficients are higher when the quadric HH income diversification measure is used with the linear term. The HH income diversification results conclude that financial institutions with higher income diversification have lower credit risk when comparing with the average credit risk in the banking sector and that asset diversification does not impact credit risk.

6.4. Effect of the Financial Crisis

The time span of the data sample used in this study starts from 2007 and ends in 2014. The outbreak of the financial crisis in 2008 could impact the results in this study, since especially financial institutions were heavily affected by the crisis. When collecting the financial statement data 2008 stood out, several financial institutions wrote off financial assets and had negative net trading income. The conditions in 2008 also affected the functional diversification and focus measures used in this study. As the trading income of the financial institutions decreased and financial assets were written down in 2008, the net interest income share of financial institution increased as income from investment banking decreased. Similarly income diversification decreased as income from investment banking decreased relative to net interest income. Thus in 2008 the functional focus and diversification measures used in this study are not fully representative, as the financial crisis had a severe impact on them. Thus in order to increase robustness the following regressions are developed in Table XIV without using data from 2008.

Table XIV – OLS Random Effects Regressions without 2008 Data

	Is Investment Banking More Risky Than Commercial Banking?			Are More Diversified Financial Institutions Less Risky?		
	CDS 5y	CDS 5y	CDS diff.	CDS 5y	CDS 5y	CDS diff.
Inc. div				-148.24 (-4.59)***	-155.46 (-2.74)***	-138.48 (-2.72)***
As. div				7.39 (0.10)	-32.54 (-0.39)	-43.84 (-0.55)
Net interest income share	-63.73 (-0.88)	-36.60 (-0.51)	-53.15 (0.82)		61.81 (0.79)	36.09 (0.51)
(Net interest income share)²		107.53 (3.95)***	89.25 (3.77)***		15.15 (0.35)	8.28 (0.22)
Loan share	75.42 (0.59)	28.22 (0.22)	-65.35 (-0.53)		122.68 (0.94)	19.49 (0.16)
Tier 1	1265.34 (2.73)***	1078.97 (2.36)**	206.94 (0.51)	1015.80 (2.34)**	1055.25 (2.33)**	219.88 (0.54)
Leverage	-4345.42 (-7.71)***	-4408.73 (-7.87)***	-3133.82 (-5.91)***	-3957.08 (-7.34)***	-4077.02 (-7.29)***	-2799.24 (-5.26)***
Cost/income.	107.92 (2.01)**	135.55 (2.55)**	108.93 (2.27)**	161.28 (4.29)***	136.00 (2.54)**	106.95 (2.22)**
LN(Size)	-106.28 (-5.28)***	-107.05 (-5.29)***	-103.90 (-5.14)***	-103.47 (-6.63)***	-85.56 (-4.00)***	-82.93 (-3.93)***
Provisions/loans	8308.44 (8.63)***	7986.85 (8.46)***	5531.87 (6.66)***	8028.40 (8.54)***	7842.49 (8.30)***	5502.01 (6.62)***
Int. exp./liabilities	-529.83 (-0.52)	-991.09 (-0.98)	203.08 (0.22)	-905.14 (-0.91)	-1042.22 (-1.04)	277.40 (0.30)
Liquid/assets	-375.93 (-0.76)	-236.70 (-0.48)	-827.80 (-1.86)*	-463.36 (-1.00)	-479.26 (-0.98)	-1008.54 (-2.27)**
Constant	1569.37 (4.90)***	1569.51 (4.89)***	1446.88 (4.57)***	1620.13 (7.53)***	1339.41 (4.15)***	1216.33 (3.83)***
N	344	344	344	344	344	344
Prob>Chi ²	0.00	0.00	0.00	0.00	0.00	0.00
R ²	0.42	0.43	0.38	0.45	0.46	0.43

* $p < 0.1$; ** $p < 0.05$; *** $p < 0.01$

When the 2008 data is excluded from the sample, the random effects results are improved with higher significance and larger coefficients. It seems that the 2008 financial crisis heavily impacted the magnitude of the previously obtained results.

When measuring the functional focus of the financial institutions with the net interest income share, the linear term is non-significant and the quadric term highly significant and positive. The coefficient of the quadric term is much larger and has higher significance than in random effects regressions with 2008 data. The results without 2008 data are confirmed with fixed effects regressions that are presented in the appendix of this study at table XVII, where coefficients are similar with similar significance as random effects without 2008 data. However, since the linear term of the net interest income share is non-significant in all of the regressions without 2008 data, no conclusions can be drawn that financial institutions focusing more in investment banking would have higher credit risk. Moreover, the positive and significant quadric net interest income share suggest that the functional diversification decreases credit risk, not that the focus of the financial institution impacts credit risk. Furthermore, as seen in Table XIV the significance of the quadric net interest income share disappears when income diversification is added to the regressions, indicating that functional diversification drives CDS spreads, not functional focus.

Since the coefficient of the income diversification is almost three times higher when the 2008 data is not used in the random effects regressions, the conclusion that functional diversification decreases the credit risk of financial institutions is further enhanced. As the 2008 data is not fully representative, the reductions in credit risk due to functional diversification seem to be much higher than previous results suggested. Using the -155 coefficient of the 5-year maturity, an increase of income diversification from the first quartile of 0.532 to the third quartile of 0.891 decreases CDS spreads by 56 basis points. When comparing with the mean 5-year CDS spread of 236 basis points (without 2008 data) the decrease has a large economic impact.

When fixed effects regressions with 2008 data were developed, the coefficients of income diversification were larger than with random effects regressions with 2008 data. Fixed effects regressions actually should not increase coefficients, leading to puzzling results. However, since the fixed effect regressions controlled for time effects, they took into account the effect of the non-representative 2008 data. Since the regressions without 2008 data have higher coefficients, it actually was logical that fixed effects regressions with 2008 data had higher coefficients than random effects regressions with 2008 data. Furthermore, in fixed effect regressions without 2008 data (presented in the appendix of this study at table XVII) the coefficients of income diversification are slightly smaller than in random effects regressions

without 2008 data (however still very significantly negative). Thus the 2008 data heavily impacted the magnitude of the results.

During the years 2011, 2012 and 2013, the European debt crisis heavily influenced the CDS spreads of the financial institutions. Several financial institutions in the sample had very high CDS spreads during those years. To make sure that the European debt crisis has no effect on the results, regressions are developed without the 2011-2013 data. Even with limiting the amount of observations, the coefficients of income diversification remained negative and highly significant, the coefficients have similar magnitude as in the main results. Regressions were also developed without data from 2011-2013 and 2008, significantly negative coefficients were gained with similar magnitude as in regressions in table XIV.

In the data sample of this study there is financial institutions from the US and Southern European countries. Especially the Southern European financial institutions have been negatively affected by the financial crisis and the Eurozone crisis during the time span of this study, which could impact the results. Also the US financial institutions could have different individual and institutional characteristics compared with European ones. To increase robustness dummy variable regressions have been developed for US and Southern European¹⁴ financial institutions with and without 2008 data and regressions have been run without US and Southern European financial institutions, with and without 2008 data. Exclusion of the US and Southern European financial institutions and the dummy variables do not impact the results, the relationship between functional diversification and CDS spreads remains significantly negative.

7. Limitations of the Study

The major limitations concerning this study mostly relate to the used data sample. The number of financial institutions used in the study is limited by the availability and liquidity of the CDS spreads of financial institutions. The sample consists of 51 financial institutions out of which 9 are from the US and the rest from Europe, making the sample Europe focused. If more liquid CDS spreads from the US could be obtained, further research could be conducted to analyse if the impact of functional diversification differs between Europe and the US. Even though only 51 financial institutions are in the sample, the used financial institutions are mostly large financial institutions that are systemically important. The financial institutions mostly

¹⁴ Spanish, Portuguese, Italian and Greece financial institutions

facing pressure to separate investment banking and commercial banking are the global systemically important financial institutions. Most of the US and European global systemically important financial institutions are represented in the sample used in this study. Conventionally using only 51 financial institutions in the sample makes generalizing results difficult. However, since most of the US and European global systemically important financial institutions are represented in the sample and they mostly face the regulatory pressure, the results of this study can be generalized to them.

Also the time span of the sample could be longer. Similarly to the number of financial institutions, the time span is restricted by the availability CDS spreads since the used CDS spread data is only available from 2007 onwards. Furthermore, the financial crisis and the European debt crisis took place during the time span of the sample. However, multiple methods have been employed to control for their effects, including the use of CDS banking sector indexes, dummy variables, time fixed effects and data without 2008 and 2011-2013 observations. The use of different robustness controls makes the results less dependent on the crises. However, results can still be effected, since the crises had such a major effect on financial institutions.

When measuring functional diversification only the mix between net interest income and non-interest income is used. However, non-interest income consists of several different income streams, like income from trading, advisory, underwriting and mutual fund services among others. Most of the different income streams can be considered as sub-functions of investments banking. By measuring the income mix in more detail with individual and more specific income streams, better conclusions could be obtained. It could be determined if there are credit risk differences between financial institutions that functionally diversify in different areas of investment banking. Providing more conclusive evidence if there are some sub-functions of investment banking that should not be combined with commercial banking.

Since in this study economies of scope are not directly measured, the negative relationship between functional diversification and credit risk cannot be interpreted as proof that economies of scope decrease credit risk in functionally diversified financial institutions. Furthermore, as agency costs or conflicts of interest are not directly measured, the negative coefficient cannot be interpreted that they do not exist, but so that the conflicts of interest and agency costs are not large enough to produce a positive coefficient. The discovered negative coefficient points more

towards establishing that diversification decreases credit risk as according to the portfolio theory (Markowitz 1952).

This study does not examine how functional diversification impacts the overall systemic credit risk in the financial institutions market, but how individual financial institutions are affected. Using models (Wagner 2010) show that functional diversification can decrease the credit risk of individual financial institutions. However, functional diversification also makes financial institutions more similar and more exposed to the same industry shocks, increasing the possibility of systemic crisis in the financial institution market (Wagner 2010). More research is called for, since it can be possible that functional diversification increases the systemic credit risk in the financial institution market and the proposition of (Wagner 2010) should be empirically tested.

8. Conclusion

The functional diversification of financial institutions has been much debated among academics and no consensus has been achieved in determining whether financial institutions should be allowed to combine investment and commercial banking. Current regulation in the US and Europe allows financial institutions to combine the two functions. However, regulators in Europe and the US are contemplating to once again separate commercial and investment banking. The purpose of this study is to examine the relationship between functional diversification and the credit risk of financial institutions from the regulative perspective. The results of this study shed light on how the credit risk of financial institutions is affected by the contemplated separation. Previous research on the effect of functional diversification on the riskiness of financial institutions has used equity based risk measures. Measuring the impact of functional diversification on credit risk is more appropriate to make regulatory conclusions, since the objective of financial institution regulation is to mostly protect the overall economy from the negative effects of financial institution defaults.

The aim of the first research question is to establish if financial institutions focusing more on investment banking have higher credit risk than those focusing more on commercial banking. When testing the relationship, the linear specification of the explanatory variable net interest income share, measuring functional focus, did not receive significant coefficients, indicating that there is no difference in credit risk between financial institutions focusing more on investment banking or commercial banking. When the non-linear specification of the functional

focus measure is used, results gained significance. However, the non-linear specification does not indicate difference in credit risk based on functional focus but potential credit risk decreases from functional diversification. Thus based on the results no difference in credit risk between financial institutions focusing more on investment banking or commercial banking was found.

The main focus of the study is to examine if financial institutions with higher functional diversification have lower credit risk in the second research question. Based on the results, functional diversification decreases the credit risk of financial institutions at the financial institution credit risk level and compared with the average credit risk in the banking sector. The results are consistent with the implementation of the portfolio theory (Markowitz 1952) to functional diversification, which suggest that diversification should decrease credit risk. The results are confirmed with multiple robustness tests by controlling for time effects, unobserved heterogeneity, endogeneity, heteroscedasticity and using alternative functional diversification measures. Robustness is also ensured by using different data samples, without data from years affected by the financial crisis and without US and Southern European financial institutions.

Based on the results, if regulators implement new regulation separating investment banking from commercial banking, the credit risk of currently functionally diversified financial institutions would increase. The separation could have serious consequences, as the probability of financial institution failures would increase. Especially since the regulators are mostly interested in breaking up the largest and most complicated functionally diversified financial institutions. As history suggests, large financial institutions failures can have grave consequences on the economy. The current economic climate does not need a new Lehman Brothers, regulators should aspire to decrease the credit risk of financial institutions, not increase it.

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10. Appendix

Table XV – Net Interest Income Share and Loan Share, Financial Institution Fixed Effects Regressions

$Net\ interest\ income\ share = \frac{INT}{TOR}$ and $Loan\ share = \frac{LOAN}{TEA}$. In the regressions the unobserved financial institution level heterogeneity is controlled for, by using dummy variables for the financial institutions. The results for the financial institution dummy variables are omitted from the table. Mainly 5 year maturity CDS spread is used in the regressions as a dependable variable, for the sake of robustness also 1 and 10 year maturities are used. CDS spreads are yearly averages of daily CDS spread observations.

	CDS 5y	CDS 5y	CDS 5y	CDS 1y	CDS 10y	CDS diff.
Net interest income share	-42.55 (-0.88)	-130.20 (-2.38)**	-127.77 (-2.33)**	-134.27 (-1.78)*	-122.07 (-2.48)**	-121.96 (-2.60)***
(Net interest income share) ²		39.41 (3.10)***	38.67 (3.03)***	58.98 (3.35)***	33.90 (2.96)***	42.63 (3.90)***
Loan share	-262.97 (-1.00)	938.07 (1.16)		-479.32 (1.33)	-405.06 (-1.72)*	-446.09 (-2.00)**
(Loan share) ²		-1127.61 (-1.64)				
Tier 1	2269.05 (4.61)***	2409.79 (4.92)***	2565.09 (5.32)***	2082.47 (3.04)***	2123.06 (4.76)***	1046.63 (2.49)**
Leverage	-7575.71 (-9.11)***	-7781.86 (-9.46)***	-7799.69 (-9.55)***	-7298.82 (-5.92)***	-5169.88 (-6.43)***	-5528.42 (-7.84)
Cost/income.	52.75 (1.56)	70.87 (2.10)**	70.69 (2.10)**	85.19 (1.83)*	68.60 (2.26)**	61.60 (2.13)**
LN(Size)	-111.46 (-1.78)*	-121.25 (-1.96)*	-100.14 (-1.71)*	-211.93 (-2.48)**	-113.79 (-2.04)**	-182.84 (-3.44)***
Provisions/loans	4820.77 (5.06)***	4327.98 (4.55)***	4706.60 (5.03)***	5855.08 (4.44)***	3758.04 (4.37)***	2659.26 (3.29)***
Int. exp./liabilities	-2773.07 (-2.59)***	-2255.43 (-2.09)**	-2584.13 (-2.45)**	-1727.64 (-1.17)	-2509.25 (-2.61)***	-1240.63 (-1.37)
Liquid/assets	219.32 (0.40)	-173.42 (0.32)	9.68 (0.02)	172.01 (0.23)	415.54 (0.84)	-833.64 (-1.78)*
N	394	394	394	390	390	394
Prob>F	0.00	0.00	0.00	0.00	0.00	0.00
R ²	0.64	0.65	0.65	0.57	0.59	0.68

* $p < 0.1$; ** $p < 0.05$; *** $p < 0.01$

Table XVI – Income and Asset Diversification, Financial Institution Fixed Effects Regressions

$INC.DIV = 1 - \left| \frac{INT-NON.INT}{TOR} \right|$ and $AS.DIV = 1 - \left| \frac{LOAN-OEA}{TEA} \right|$. In the regressions the unobserved financial institution level heterogeneity is controlled for, by using dummy variables for the financial institutions. The results for the financial institution dummy variables are omitted from the table. Net interest income share and loan share ratios are used as control variables. Mainly 5 year maturity CDS spread is used in the regressions as a dependable variable, for the sake of robustness also 1 and 10 year maturities are used. CDS spreads are yearly averages of daily CDS spread observations.

	CDS 5y	CDS 5y	CDS 5y	CDS 1y	CDS 10y	CDS diff.
Inc. div	-71.92 (-3.13)***	-77.30 (-3.34)***	-60.65 (-1.13)	-118.77 (-3.72)***	-70.27 (-3.38)***	-79.10 (-3.97)***
As. div	272.07 (2.00)**	290.46 (2.13)**	286.30 (2.09)**	374.12 (1.99)**	308.14 (2.52)**	271.17 (2.31)**
Net interest income share		-75.84 (-1.58)	-92.26 (-1.36)	-60.21 (-0.91)	-84.19 (-1.95)*	-67.97 (-1.64)
(Net interest income share) ²			10.12 (0.34)			
Tier 1	2414.70 (5.02)***	2482.04 (5.15)***	2485.91 (5.15)***	2158.19 (3.21)***	2182.77 (4.99)***	1114.50 (2.69)***
Leverage	-7594.07 (-9.33)***	-7597.14 (-9.35)***	-7658.82 (-9.20)***	-7160.31 (-5.89)***	-5164.30 (-6.53)***	-5481.08 (-7.84)***
Cost/income.	21.75 (1.07)	63.08 (1.90)*	66.40 (1.92)*	77.92 (1.70)*	67.52 (2.27)**	57.27 (2.01)**
LN(Size)	-111.16 (-1.89)*	-112.17 (-1.91)*	-112.58 (-1.92)*	-191.07 (-2.36)**	-96.42 (-1.83)*	-159.64 (-3.16)***
Provisions/loans	4544.24 (4.85)***	4375.45 (4.65)***	4376.82 (4.64)***	5664.27 (4.31)***	6310.42 (4.22)***	2587.34 (3.19)***
Int. exp./liabilities	-2231.35 (-2.09)**	-2159.24 (-2.02)**	-2170.75 (-2.03)**	-1080.73 (-0.73)	-1989.60 (-2.06)**	-761.16 (-2.25)**
Liquid/assets	90.68 (0.17)	78.95 (0.15)	78.20 (0.15)	-19.72 (-0.03)	256.44 (0.54)	-1020.23 (-2.25)**
N	394	394	394	390	390	394
Prob>F	0.00	0.00	0.00	0.00	0.00	0.00
R ²	0.65	0.65	0.65	0.57	0.60	0.68

* $p < 0.1$; ** $p < 0.05$; *** $p < 0.01$

**Table XVII – OLS Financial Institution and Year Fixed Effects Regressions
Without 2008 Data**

	Is Investment Banking More Risky Than Commercial Banking?			Are More Diversified Financial Institutions Less Risky?		
	CDS 5y	CDS 5y	CDS diff.	CDS 5y	CDS 5y	CDS diff.
Inc. div				-128.69 (-4.26)***	-138.43 (-4.21)***	-140.09 (-4.27)***
As. div				99.25 (0.77)	127.92 (0.92)	130.25 (0.94)
Net interest income share	-70.29 (-0.94)	0.66 (0.01)	-21.45 (-0.29)		59.15 (0.75)	40.01 (0.50)
(Net interest income share)²		106.70 (4.76)***	104.54 (4.65)***			
Loan share	87.26 (0.17)	-9.78 (-0.04)	-56.34 (-0.23)		143.11 (0.55)	95.82 (0.37)
Tier 1	87.87 (0.17)	158.78 (-0.31)	261.49 (-0.51)	169.51 (0.34)	135.84 (0.26)	28.92 (0.06)
Leverage	-6065.29 (-7.79)***	6155.62 (-8.20)***	-5712.69 (-7.60)***	-5868.00 (-7.82)***	-5837.78 (-7.65)***	-5394.36 (-7.07)***
Cost/income.	53.95 (0.97)	51.02 (0.95)	60.65 (1.13)	81.27 (2.55)**	48.73 (0.90)	58.24 (1.08)
LN(Size)	-214.12 (-3.54)***	-241.36 (-4.13)***	-208.57 (-3.56)***	-218.77 (-3.95)***	-210.23 (-3.58)***	-177.95 (-3.03)***
Provisions/loans	4126.17 (4.43)***	4439.55 (4.93)***	3439.70 (3.81)***	4199.65 (4.67)***	4329.86 (4.74)***	3338.22 (3.66)***
Int. exp./liabilities	3460.39 (2.37)**	2869.13 (2.03)**	2436.12 (1.72)*	3321.57 (2.38)**	3645.19 (2.53)**	3203.86 (2.22)**
Liquid/assets	-1639.34 (-3.11)***	-1501.25 (-2.96)***	-1747.61 (-3.43)***	-1498.64 (-3.05)***	-1561.81 (-3.05)***	-1804.37 (-3.52)***
N	344	344	344	344	344	344
Prob>F	0.00	0.00	0.00	0.00	0.00	0.00
R ²	0.79	0.80	0.76	0.80	0.80	0.76

* $p < 0.1$; ** $p < 0.05$; *** $p < 0.01$